OR Forum—The Cost of Latency in High-Frequency Trading

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Modern electronic markets have been characterized by a relentless drive toward faster decision making. Significant technological investments have led to dramatic improvements in latency, the delay between a trading decision and the resulting trade execution. We describe a theoretical model for the quantitative valuation of latency. Our model measures the trading frictions created by the presence of latency, by considering the optimal execution problem of a representative investor. Via a dynamic programming analysis, our model provides a closed-form expression for the cost of latency in terms of well-known parameters of the underlying asset. We implement our model by estimating the latency cost incurred by trading on a human time scale. Examining NYSE common stocks from 1995 to 2005 shows that median latency cost across our sample roughly tripled during this time period. Furthermore, using the same data set, we compute a measure of implied latency and conclude that the median implied latency decreased by approximately two orders of magnitude. Empirically calibrated, our model suggests that the reduction in cost achieved by going from trading on a human time scale to a low latency time scale is comparable with other execution costs faced by the most cost efficient institutional investors, and it is consistent with the rents that are extracted by ultra-low latency agents, such as providers of automated execution services or high frequency traders.

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1. Introduction

In the past decade, electronic markets have become pervasive. Technological advances in these markets have led to dramatic improvements in latency, or the delay between a trading decision and the resulting trade execution. In the past 30 years, the time scale over which a trade is processed has gone from minutes to milliseconds—“low latency” in a contemporary electronic market would be qualified as under 10 milliseconds, “ultra-low latency” as under one millisecond. This change represents a dramatic reduction by five orders of magnitude. To put this in perspective, human reaction time is thought to be in the hundreds of milliseconds.

One factor behind this trend has been competition between exchanges, as one mechanism for differentiation between exchanges is latency. This competition is driven by a significant demand among a class of investors, sometimes called “high-frequency” traders, for low latency trade execution. High-frequency traders are thought to account for more than half of all U.S. equity trades. They expend significant resources to develop algorithms and systems that are able to trade quickly. For example, on the time scale of milliseconds, the speed of light can become a binding constraint on the delay in communications. Hence, traders seeking low latency will “co-locate,” or house their computers in the same facility as the exchange, to eliminate delays due to a lack of physical proximity. This co-location comes at a significant expense; however, it has been stated that a one-millisecond advantage can be worth $100 million to a major brokerage firm.

There has been much discussion of the importance of latency among various market participants, regulators, and academics. Despite the significant amount of recent interest, however, latency remains poorly understood from a theoretical perspective. For example, how does latency relate to transaction costs? Is latency relevant only to investors with short time horizons, such as high-frequency traders, or does latency also affect long-term investors such as pension funds and mutual funds? Many of these important questions have been considered in anecdotal or ad hoc discussions. Our goal here is to provide a framework for quantitative analysis of these issues.

In particular, we wish to understand the benefit to a single trader in the marketplace of lowering their latency, while holding everything else fixed. This is a different question than understanding the social costs of latency, i.e., whether in equilibrium the collective marketplace is better or worse off given lower latency. One might imagine,
for example, that the benefit to individual agents of lower latency could diminish in an equilibrium setting. Equilibrium or welfare analysis of low latency trading is a complex question with important policy and regulatory implications. We believe that understanding the single-agent effects of low latency trading, however, is an important first step that will inform our ultimate understanding of collective effects.

The cost that a trader bears due to latency can take many different forms, depending on the precise trading strategy. However, we can identify a number of broad themes, sometimes overlapping, concerning why the ability to trade with low latency might be valuable to an investor:

1. **Contemporaneous decision making.** A trader with significant latency will be making trading decisions based on information that is stale.

   For example, consider an automated trader implementing a market-making strategy in an electronic limit order book. The trader will maintain active limit orders to buy and sell. The prices at which the trader is willing to buy or sell will naturally depend on, say, the limit orders submitted by other investors, the price of the asset on other exchanges, the price of related assets, overall market factors, etc. If the trader cannot update his orders in a timely fashion in response to new information, he might end up trading at disadvantageous prices.

2. **Comparative advantage/disadvantage.** The ability to trade with low latency in absolute terms might not be as important as the ability to trade with low relative latency, that is, as compared to competitors.

   For example, consider a program trader implementing an index arbitrage strategy, seeking to profit on the difference between an index and its underlying components. There might be many market participants pursuing such strategies and identifying the same discrepancies. The challenge for the trader is to be able to act in the marketplace to exploit a discrepancy before a price correction takes place, i.e., before competitors are able to act. The means having a low relative latency.

3. **Time priority rules.** Many modern markets treat orders differentially based on the time of arrival, and favor earlier orders.

   For example, in an electronic limit order book, the limit orders on each side of the market are prioritized in a particular way. When a market order to buy arrives, it is matched against the limit orders to sell according to their priorities. Priority is first determined by price, i.e., limit orders with more lower prices receive higher priority. In many markets, however, prices are mandated to be discrete with a minimum tick size. In these markets, there might be multiple limit orders at the same price, which are then prioritized according to the time of their arrival. While a trader can always increase the priority of his orders by decreasing price, this comes at an obvious cost. If a trader can submit orders in a faster fashion, however, he can increase priority while maintaining the same price. Higher priority can be valuable for two reasons: First, higher priority orders have a higher likelihood of execution over any given time horizon. To the extent that investors submitting limit orders have a desire to trade, and to trade sooner rather than later, this is desirable. Second, higher priority orders at the same price level experience less adverse selection (see, e.g., Glosten 1994, Sandás 2001). Hence, all things being equal, an investor who submits orders with lower latency will benefit from higher priority than if that investor had higher latency. This can be particularly important (in that a small improvement in latency can result in a significant difference in priority) when an existing quote is about to change. For example, consider the situation where a stock price is about to move up because of trades or cancellations at the best offered price. One might expect the bid price to rise as well, there will be a race among traders reacting to the same order book events to establish time priority at the new bid.

In this paper, we will quantify the cost of latency due to the first effect, a lack of contemporaneous decision making. We do not consider effects of latency that arise from strategic considerations, or from time priority rules or price discreteness. It is an open question as to whether the other effects are more or less significant than the first, and their relative importance might depend on the particular investor and their trading strategy. Our analysis does not speak to this point. However, in what follows we will demonstrate that, by itself, the lack of contemporaneous decision making can induce trading costs that are of the same order of magnitude as other execution costs faced by large investors, and hence cannot be neglected.

Furthermore, the importance of contemporaneous decision making will certainly vary from investor to investor. We will focus on an aspect of this that is universal, however, which is the importance of timely information for the execution of contingent orders. A contingent order, such as a limit order in an electronic limit order book or a resting order in a dark pool, presents the possibility of uncertain execution over an interval of time in exchange for price improvement relative to a market order, which executes immediately and with certainty. Specifically, when an investor employs a contingent order, the investor might be exposed to the realization of new information (for example, in the form of price movements, news, etc.) over the lifespan of the order. Latency, which prevents the investor from continuously and instantaneously accessing the market so as to update the order, can thus adversely impact the investor.

As a broad proxy for understanding the importance of latency in contingent order execution, we consider the effects of latency in an extremely simple yet fundamental trade execution problem: that of a risk-neutral investor who wishes to sell 1 share of stock (i.e., an atomic unit) over a fixed, short time horizon (i.e., seconds) in a limit order book, and must decide between market orders and limit orders. Our problem formulation is reminiscent of barrier-diffusion models for limit order execution (e.g., Harris 1998). It captures the fundamental cost
of immediacy of trading (e.g., Grossman and Miller 1988, Chacko et al. 2008), that is, the premium due to a patient liquidity supplier (who submits limit orders) relative to an impatient demander of liquidity (who submits market orders). While this problem is quite stylized, we will argue that it is broadly relevant since, at some level, all investors make such a choice of immediacy. For example, it might not seem at first glance that our execution problem is relevant for a pension fund that trades large blocks of stock over multiple days. However, the execution of a block trade via algorithmic trading involves the division of a large “parent” order into many atomic orders over the course of a day; each of these atomic “child” orders can be executed as limit orders or as market orders.

In our problem, in the absence of latency, the optimal strategy of the seller is a “pegging” strategy: the seller maintains a limit order at a constant spread above the bid price at any instant in time. We consider this case as a benchmark. In the presence of latency, the seller can no longer maintain continuous contact with the market so as to track the bid price in the market. The seller is forced to deviate from the benchmark policy in order to take into account the uncertainty introduced by the latency delay by incorporating a safety margin and lowering his limit order prices. The friction introduced by latency thus results in a loss of value to the seller. We will establish the difference in value to the seller between the case with latency and the benchmark case via dynamic programming arguments, and thus provide a quantification of the effects of latency.

The contributions of this paper are as follows.

• We mathematically quantify the cost of latency. The trading problem we consider (deciding between limit and market orders) is faced by all large investors in modern equity markets, either directly (e.g., high-frequency traders) or indirectly (e.g., pension funds who execute large trades via providers of automated execution services). Our analysis suggests that latency impacts all these market participants and that, all else being equal, the ability to trade with low latency results in quantifiably lower transaction costs. Furthermore, when calibrated with market data, the latency cost we measure can be significant. It is of the same order of magnitude as other trading costs (e.g., commissions, exchange fees, etc.) faced by the most cost-efficient large investors. Moreover, it is consistent with the rents that are extracted by agents who have made the requisite technological investments to trade with ultra-low latency. For example, the latency cost of our model is comparable to the execution commissions charged by providers that offer algorithmic trade execution services on an agency basis. It is also comparable to the reported profits of high-frequency traders.

To our knowledge, our model is the first to provide a quantification of the costs of latency in trade execution.

• We provide a closed-form expression for the cost of latency as a function of well-known parameters of the asset price process.

The cost of latency in our model can be computed numerically via dynamic programming. However, in the regime of greatest interest, where the latency is close to zero, we provide a closed-form asymptotic expression. In particular, define the latency cost associated with an asset as the costs incurred due to latency as a fraction of the overall cost of immediacy (the premium paid to a patient liquidity supplier by an impatient demander of liquidity). Given a latency of \( \Delta t \), a price volatility of \( \sigma \), and a bid-offer spread of \( \delta \), the latency cost takes the form

\[
\frac{\sigma \sqrt{\Delta t}}{\delta} \sqrt{\log \frac{\delta^2}{2\pi \sigma^2 \Delta t}}
\]  
(1)

as \( \Delta t \to 0 \).

• Our method can provide qualitative insight into the importance of latency.

From (1), it is clear that the latency cost is an increasing function of the ratio of the standard deviation of prices over the latency interval (i.e., \( \sigma \sqrt{\Delta t} \)) to the bid-offer spread. Latency has a more important role when trading assets that are either more volatile (\( \sigma \) large) or, alternatively, more liquid (\( \delta \) small). Furthermore, as the latency approaches 0, the marginal benefit of latency reduction is increasing.

• We empirically demonstrate that latency cost incurred by trading on a human time scale has dramatically increased for U.S. equities and the implied latency of a representative trader in this market decreased by approximately two orders of magnitude.

We consider the cost due to the latency of trading on the time scale of human interaction. Using the data set of Aït-Sahalia and Yu (2009), we estimate the latency cost of NYSE common stocks over the 1995–2005 period. We show that the median latency cost roughly tripled in this time. This coincides with a period of decreasing tick sizes and increasing algorithmic and high-frequency trading activity (Hendershott et al. 2011).

An alternative perspective is to consider a hypothetical investor who fixes a target level of cost due to latency, relative to the overall cost-of-immediacy. The representative trader maintains this target over time through continual technological upgrades to lower levels of latency. We determine the requisite level of implied latency for such a trader, over time and across the aggregate market. Using the same data set, we observe that the median implied latency decreased by approximately two orders of magnitude over this time frame.

The rest of this paper is organized as follows. In §1.1, we review the related literature. In §2, as a starting point, we present a stylized, continuous-time trade execution problem in the absence of latency. We develop a variation of the model with latency in §3. In §4, we provide a mathematical analysis of the optimal policy for our problem. By contrasting the results in the presence and absence of latency, we are able to quantitatively assess the cost of latency. In §5, we consider some empirical applications of the model. Finally, in §6 we conclude and discuss some future directions. Supplemental material to this paper is available at http://dx.doi.org/10.1287/opre.2013.1165.
1.1. Related Literature

There has been a significant empirical literature studying, broadly speaking, the effects of improvements in trading technology. Closest to the aspect we consider is the work of Easley et al. (2013). They empirically test the hypothesis that latency affects asset prices and liquidity by examining the time period around an upgrade to the New York Stock Exchange technological infrastructure that reduced latency. Hendershott et al. (2011) explore the more general, overall effects of algorithmic and high-frequency trading. Hasbrouck and Saar (2009) provide different evidence of changes in investor trading strategies that might be a result of improved technology. In subsequent work, they further consider the impact of measurements of low latency on market quality (Hasbrouck and Saar 2010). Hendershott and Riordan (2009) analyze the impact of algorithmic trading on the price formation process using a data set from Deutsche Börse and conclude that algorithmic trading assists in the efficient price discovery without increasing the volatility. Kirilenko et al. (2010) consider the impact of high frequency trading on the “flash crash” of 2010, while Brogaard (2010) more broadly examines the impact of high-frequency traders on market quality.

On the theoretical front, Cespa and Foucault (2008) consider a rational expectations equilibrium between investors with different access to past transaction data. Some investors observe transactions in real time, while others observe transactions only with a delay. This model of latency focuses on latency of the price ticker of past transactions, as opposed to latency in execution, which we consider here. Moreover, the goals of the two models differ significantly; Cespa and Foucault (2008) seek to build intuition regarding the equilibrium welfare implications of differential access to information via a structural model; we, on the other hand, seek a reduced form model that can be used to directly estimate the value of execution latency in a particular real-world instance, given readily available data. Also related is the work of Ready (1999) and Stoll and Schenzler (2006), who consider the ability of intermediaries (e.g., specialists or dealers) to delay customer orders for their own benefit, thus creating a “free option” in the presence of execution latency. Cohen and Szpruch (2012) show that latency arbitrage exists between two traders with different speeds of trading in the presence of a limit order book. Finally, Cvitanić and Kirilenko (2010) and Jarrow and Protter (2011) consider the effect of high-frequency traders on asset prices.

The trade execution problem we consider is that of an investor who wishes to sell a single share and must decide between market and limit orders. This problem has been considered by many others (e.g., Angel 1994, Harris 1998, Lo et al. 2002). Our formulation is similar to the class of barrier-diffusion models considered by these authors; Hasbrouck (2007) provides a good account of this line of work. For a broad survey on limit order markets, see Parlour and Seppi (2008). In our model, the inability to trade continuously gives a limit order an option-like quality that relates execution cost, order duration, and asset volatility. This idea goes as far back as the work of Copeland and Galai (1983). Closely related is the concept of the cost of immediacy, or the premium paid by a liquidity demander via a market order to a liquidity supplier who posts a limit order. Grossman and Miller (1988) and Chacko et al. (2008) develop theoretical explanations of the cost of immediacy. For empirical evidence of the demand for immediacy in capital markets, see Bacidore et al. (2003) and Werner (2003).

Finally, also related is work on the discrete-time hedging of contingent claims with or without transaction costs (e.g., Boyle and Emanuel 1980, Leland 1985, Bertsimas and Werner (2003)). This literature addresses a different problem and draws different conclusions than our paper, however both relate to implications of a lack of continuous access to the market.

2. A Stylized Execution Model Without Latency

Our goal is to understand the impact on the trade execution of latency. To this end, we will first describe a trade execution problem in the absence of latency. In §3, we will revisit this model in the presence of latency to understand the resulting trade friction that is introduced. The spirit of our model is to consider an investor who wants to trade, but at a price that depends on an informational process that evolves stochastically and must be monitored continuously. We could directly consider such an abstract model of investor behavior. Instead, however, we will motivate the informational dependence of the trader through a specific optimal execution problem.

Consider the following stylized execution problem of an uninformed trader who must sell exactly one share of a stock over a time horizon [0, T]. At any time \( t \in [0, T] \), the trader can take one of two actions:

1. The trader can submit a market order to sell. This order will execute at the best bid price at time \( t \), denoted by \( S_t \). We assume that the bid price evolves according to

\[
S_t = S_0 + \sigma B_t,
\]

where the process \((B_t)_{t \in [0, T]}\) is a standard Brownian motion and \( \sigma > 0 \) is an (additive) volatility parameter. Here, the choice of Brownian motion is made for simplicity; our model can be extended to the more general class of Markovian martingales, as discussed in §4.4.

2. The trader can choose to submit a limit order to sell. In this case, the trader must also decide the limit price associated with the order, which we denote by \( L_t \).

Once the trader sells one share, he exits the market. If the trader is not able to sell 1 share before time \( T \), however, we assume that he is forced sell via a market order at time \( T \) and therefore receives \( S_T \). Here, we imagine the time horizon \( T \) to be small, on the order of the typical trade execution time (i.e., seconds).
2.1. Limit Order Execution
It remains to describe the execution of limit orders. In our setting, a limit order can execute in one of the following two ways:

1. We assume that there are impatient buyers who arrive to the market according to a Poisson process with rate \( \mu \). Denote by \((N(t))_{t \geq 0}\) the cumulative arrival process for impatient buyers. Each impatient buyer seeks to buy a single share. An arriving impatient buyer at time \( t \) has a reservation price \( S_t + z_t \), expressed as a premium \( z_t \geq 0 \) above the bid price \( S_t \) that the buyer is willing to forgo in order to achieve immediate execution. We assume that the premium \( z_t \) is independent and identically distributed with cumulative distribution function \( F: \mathbb{R}_+ \rightarrow [0, 1] \). In this setting, the instantaneous arrival rate of impatient buyers at time \( t \) willing to pay a limit order price of \( L_t \) is given by

\[
\lambda(u_t) = \mu(1 - F(u_t)),
\]

where \( u_t \triangleq L_t - S_t \) is the instantaneous price premium of the limit order. In what follows, we will be particularly interested in the special case where

\[
\lambda(u_t) = \begin{cases} \mu & \text{if } u_t \leq \delta, \\ 0 & \text{otherwise}. \end{cases}
\]

Here, we assume that every impatient buyer is willing to pay a price premium of at most \( \delta > 0 \). We assume that \( \delta \) will be specific to the security and fixed for the trading horizon. We will discuss the extension to the general case (3) in §4.4.

Given (4), an impatient buyer is willing to buy 1 share at a fixed premium \( \delta > 0 \) to the bid price at the time of their arrival. Hence, if a buyer arrives at time \( \tau \in [0, T] \), and the trader has placed a limit order with price \( L_{\tau} \), the limit order will execute if \( L_{\tau} \leq S_{\tau} + \delta \).

2. Alternatively, a limit order will also execute at time \( \tau \) if the bid price crosses the limit order price, i.e., \( S_{\tau} \geq L_{\tau} \).

The execution of limit orders in the model is illustrated in Figure 1.

The limit order execution dynamics above can also be economically interpreted in the spirit of the non-informational trade model of Roll (1984). In particular, imagine that the asset has a fundamental value \( V_t \) at time \( t \), and that \( V_t \) evolves exogenously according to the additive random walk

\[
V_t = V_0 + \sigma B_t.
\]

If all investors observe this underlying value process and are symmetrically informed, competitive market makers will always be willing to sell shares at a price of \( \delta/2 \) above the fundamental value or buy shares at a spread of \( \delta/2 \) below the fundamental value. Here, the quantity \( \delta \) captures the per-share operating costs of trade to the market makers. The liquidating trader can thus sell at the bid price

\[
S_t = V_t - \delta/2 \quad \text{at any time } t.
\]

We assume that all other traders in the market are impatient and that these traders arrive according to the Poisson dynamics described above. An arriving impatient buyer will choose to purchase from the liquidating trader only at a price lower than that provided by the market makers, i.e., only below the price of \( V_t + \delta/2 = S_t + \delta \). In this way, we can interpret the parameter \( \delta \) as the prevailing bid-offer spread; that is, the bid-offer spread in the absence of the liquidating trader.

2.2. Optimal Solution
Let \( P \) denote the random variable associated with the sale price. We assume the trader is risk-neutral and seeks to maximize the expected sale price. Equivalently, we assume the trader seeks to solve the optimization problem

\[
\tilde{h}_0 \triangleq \maximize \mathbb{E}[P] - S_0.
\]

Here, the maximization is over policies of market orders and limit orders that are nonanticipating, i.e., policies adapted to the filtration generated by the underlying market primitives, \((B_t, N_t)_{t \in [0, T]}\). This objective is equivalent to minimizing implementation shortfall (Perold 1988).

Note that while this stylized problem might seem quite simplified, it seeks to answer a fundamental question: at the level of an atomic unit of stock and over a short time horizon, how should a risk-neutral investor choose between limit orders and market orders? This problem is a central ingredient in more sophisticated optimal execution problems involving risk-averse investors selling large quantities over longer time horizons. This is because in a typical algorithmic trading setting, a large “parent” order will be scheduled across time into many very small “child” orders. Each of these child orders needs to be executed optimally. Since each child order is small and since there are many such child orders, it is reasonable to view the investor as risk-neutral with respect to each child order.
The following lemma characterizes a simple strategy that is optimal for the execution problem we have described.

**Lemma 1.** An optimal strategy is to employ only limit orders at times \( t \in [0, T] \), with limit price \( L_t = S_t + \delta \). In other words, the limit order price is "pegged" at a constant premium \( \delta \) above the bid price. This pegging strategy achieves the optimal value

\[
\tilde{h}_0 = \delta(1 - e^{-\mu T}).
\]

**Proof.** Consider a trader using an arbitrary strategy, and denote by \( \tau \in [0, T] \) the (random) time at which the trader sells the share, and by \( \tau_l \in [0, \infty) \) the time at which the first impatient buyer arrives to the market. Let \( \mathcal{E} \) be the event that the trader sells via a limit order to an impatient buyer at the price \( L_\tau \). Then, under the event \( \mathcal{E} \), the trader sells at the bid price \( S_\tau \). Then the sale price \( P \) can be written as

\[
P = S_\mathcal{E} + L_\mathcal{E} + (S_\tau + \delta) \mathbb{1}_{\mathcal{E}} \leq S_\tau + \delta \mathbb{1}_{\{\tau_l < T\}}.
\]

Here, for the first inequality, we used the fact that an impatient buyer will buy only at time \( \tau \) \( \leq S_\tau + \delta \), and, for the second inequality, we used the fact that the event \( \mathcal{E} \) can only occur if an impatient buyer arrives in the time interval \( [0, \tau] \). Denote by \( \tilde{h}_0 \) the value under an optimal strategy. Using the fact that \( \tau \) is a bounded stopping time and the fact that \( S_\tau \) is a martingale, by the optional sampling theorem,

\[
\tilde{h}_0 \leq \mathbb{E}[P] - S_0 \leq \mathbb{E}[S_\tau + \delta \mathbb{1}_{\{\tau_l < T\}}] - S_0 = \delta \mathbb{P}(\tau_l < T) = \delta(1 - e^{-\mu T}).
\]

On the other hand, the hypothesized strategy results in equality in (7). Thus, the result follows. \( \square \)

The optimal pegging strategy suggested by Lemma 1 is illustrated in Figure 2. This policy can be interpreted intuitively as follows: since the trader is risk-neutral and the bid-price process is a martingale, the trader is indifferent as to whether to immediately sell the share, and by \( \tau_l \in [0, \infty) \) the time at which the first impatient buyer arrives to the market. Let \( \mathcal{E} \) be the event that the trader sells via a limit order to an impatient buyer at the price \( L_\tau \). Then, under the event \( \mathcal{E} \), the trader sells at the bid price \( S_\tau \). Then the sale price \( P \) can be written as

\[
P = S_\mathcal{E} + L_\mathcal{E} + (S_\tau + \delta) \mathbb{1}_{\mathcal{E}} \leq S_\tau + \delta \mathbb{1}_{\{\tau_l < T\}}.
\]

Here, for the first inequality, we used the fact that an impatient buyer will buy only at time \( \tau \) \( \leq S_\tau + \delta \), and, for the second inequality, we used the fact that the event \( \mathcal{E} \) can only occur if an impatient buyer arrives in the time interval \( [0, \tau] \). Denote by \( \tilde{h}_0 \) the value under an optimal strategy. Using the fact that \( \tau \) is a bounded stopping time and the fact that \( S_\tau \) is a martingale, by the optional sampling theorem,

\[
\tilde{h}_0 \leq \mathbb{E}[P] - S_0 \leq \mathbb{E}[S_\tau + \delta \mathbb{1}_{\{\tau_l < T\}}] - S_0 = \delta \mathbb{P}(\tau_l < T) = \delta(1 - e^{-\mu T}).
\]

3. **A Model for Latency**

The optimal policy for the stylized execution problem of §2 relied on the ability of a trader to continuously track an informational process, namely, the bid price in the market, and to update his order as the process evolves. Here, we will consider a variation of that problem where the trader is unable to continuously participate in the market but faces a fixed latency \( \Delta t > 0 \).\(^8\) We are interested in quantifying the cost of this latency by comparing the expected payoff in this model to that in the stylized model without latency. Note that the model at hand is quite basic with regard to some primitives (e.g., the stylized process describing the evolution of bid prices); we will discuss a number of tractable extensions in §4.4, including more complicated models of the bid-price process and of limit order execution.

In general, latency that a trader experiences can take many forms. Minimally, for example, there is the delay of the data feeds that deliver market price information to the trader. There is the delay of the trader’s own decision making. Finally, there is the delay of the trader’s resulting order reaching the marketplace. We assume that the trader makes decisions instantaneously—we will see that this is reasonable since the optimal decision rule for the trader will take...
a very simple form. Furthermore, from the trader’s perspective, the roundtrip delay (the total delay for an order to be processed by an exchange and reflected in the data feeds observed by the trader) cannot be decomposed into a delay to the exchange and a delay from the exchange. Hence, without loss of generality, we will assume that the trader is able to observe market price information with no delay or latency, but that the trader’s orders experience a latency $\Delta t$ before they are processed by the exchange. This latency is meant to capture, for example, networking or routing delays that are specific to the trader, and that might be reduced through colocation or additional investment in networking technology.

In our latency model, we consider an investor who maintains a limit order to sell one share over the time horizon $[0, T]$ (the possibility of market orders will be discussed shortly), so that once the limit order is executed, the investor immediately exits the market. The time horizon $[0, T]$ is divided into $n$ slots each of length $\Delta t$, i.e., $T = n\Delta t$. For each $i \in \{0, 1, \ldots, n\}$, define $T_i \triangleq i\Delta t$.

At each time $T_i$, based on all information observed thus far, we assume that the trader can instantaneously decide to update the limit order with a new price $l_i$. Due to a latency of $\Delta t$, the updated price does not reach the market and take effect until the beginning of the next time slot, i.e., $T_{i+1}$. This limit order price remains active until time $T_{i+2}$, at which point it is superseded by the next price $l_{i+1}$. This sequence of events is illustrated in Figure 3. Between the time $T_i$, when the price $l_i$ is decided, and the time $T_{i+1}$, when the updated order reaches the market, the following events can occur:

- $\xi_{i}^{(1)}$: An impatient buyer arrives in the time interval $(T_i, T_{i+1})$ and $l_{i-1} \leq S_{T_i} + \delta$, i.e., the prior limit price $l_{i-1}$, which is active at that time, is within a margin $\delta$ of the bid price at the start of the interval. In this case, the limit order executes at the price $l_{i-1}$, and the investor leaves the market. Note that the updated limit price $l_i$ never takes effect.

- $\xi_{i}^{(2)}$: Otherwise, if $S_{T_{i+1}} \geq l_i$, i.e., the bid price has crossed the order price $l_i$ at the instant the order reaches the market, then the order immediately executes at price $S_{T_{i+1}}$.

- $\xi_{i}^{(3)}$: Otherwise, the limit order price $l_i$ is active over the time interval $[T_{i+1}, T_{i+2})$.

To consider the possibility of market orders, we allow the limit price $l_i = -\infty$. By picking this price, the trader can guarantee that the bid price at time $T_{i+1}$ will cross the order price, i.e., $S_{T_{i+1}} \geq l_i$, with probability 1. Thus, the choice of $l_i = -\infty$ corresponds to a certain execution at the bid price $S_{T_{i+1}}$, i.e., a market order. Similarly, the trader can make the decision at time $T_i$ not to trade by setting $l_i = \infty$. As in the model of §2, if the investor has been unable to sell the share by the end of the time horizon $T$, the investor is forced to sell via a “clean-up” trade, i.e., a market order at time $T$. This is accomplished by enforcing the constraint that $l_{n-1} = -\infty$, which we will assume implicitly in what follows.

As before, if $P$ is the random variable associated with the sale price, the trader is risk-neutral and seeks to solve the optimization problem

$$
\hat{h}_0(\Delta t) \triangleq \maximize_{l_0, \ldots, l_{n-1}} E[P] - S_0.
$$

Here, the maximization is over the choice of limit order prices $(l_0, l_1, \ldots, l_{n-1})$. We assume that the price decisions are non-anticipating, i.e., each $l_i$ is adapted to the filtration generated by the bid price process and the arrival of impatient buyers up to and including time $T_i$. Our goal is to analyze $h_0(\Delta t)$, which is the value under an optimal trading strategy when the latency is $\Delta t$.

Note that, as compared to the model of §2, our present model with latency differs in two ways: First, the trader makes decisions at the beginning of discrete-time intervals of length $\Delta t$, as opposed to continuously. Second, the orders of the trader incur a latency or delay of length $\Delta t$ before they reach the marketplace. We are interested in studying the impact of the latter feature, latency, and we adopt the former feature, discrete-time decision making, so as to admit a tractable dynamic programming analysis. In §4.3, however, we will see that in the low latency regime in which we are most interested, the discrete-time nature of our model has a negligible impact.

### 4. Analysis

In this section, we solve for the optimal policy for the trader in the latency model of §3. This problem can be solved via a dynamic programming decomposition that is presented in §4.1. While the exact dynamic programming solution can be computed numerically, in §4.2 we will present an asymptotic analysis that provides a closed-form analytic expression for the cost of latency in the low latency regime, where $\Delta t \to 0$. In §4.3, we will consider the implications of the discrete-time nature of our latency model. Finally, in §4.4, we will discuss a number of extensions of our latency model.
4.1. Dynamic Programming Decomposition

The standard approach to solving the optimal control problem (8) is to employ dynamic programming arguments. Supplemental material to this paper is available at http://dx.doi.org/10.1287/opre.2013.1165. In Appendix A of the electronic companion, we formally derive the optimal control policy using these methods. To focus on the high-level picture, however, for the moment we will be content with summarizing those results.

In particular, assume a fixed latency of $\Delta t$. For each decision time $T_i$ with $0 \leq i < n$, define $h_i$ to be the event that the trader’s limit order remains unfulfilled prior to time $T_i$, i.e., none of the orders submitted at prices $l_1, \ldots, l_{i-1}$ are executed. Note that if the event $h_i$ does not hold, then the limit order price $l_i$ to be decided at time $T_i$ is irrelevant. This is because by the time that order arrives to the market, the trader would have already sold a share. Define the quantity

$$h_i \triangleq \text{maximize } E[P \mid S_{T_i}, h_i] - S_{T_i}. \tag{9}$$

Note that $h_0 = h_0(\Delta t)$, where $h_0(\Delta t)$ is defined in (8), and thus our notation is consistent. More generally, for $i > 0$, we can interpret $h_i$ to be the trader’s expected payoff at time $T_i$ relative to the current bid price $S_i$ under the optimal policy, the order does not get filled prior to time $T_{i+1}$. Thus, $h_i$ can be interpreted as a continuation value in the dynamic programming context.

The continuation values $\{h_i\}$ quantify the remaining value for a trader at each time period if his order remains unfulfilled. Given the continuation values, at each time $T_i$, the investor can make an optimal decision as to the limit order price $l_i$ by balancing the benefits of execution in the time slot $[T_{i+1}, T_{i+2})$ with the value $h_{i+1}$ that will be obtained if the order is not executed. Moreover, the optimal decisions and continuation values can be jointly computed via backward induction of a Bellman equation. This result is captured in the following theorem. The proof, which is provided in Appendix A of the electronic companion, follows from formal dynamic programming arguments.

**Theorem 1.** Suppose $\{h_i\}$ satisfy, for $0 \leq i < n - 1$,

$$h_i = \max_{u_i} \left\{ \mu \Delta t \left[ u_i \left( \Phi \left( \frac{u_i - \delta}{\sigma \sqrt{\Delta t}} \right) - \Phi \left( \frac{u_i - \delta}{\sigma \sqrt{\Delta t}} \right) \right) + \sigma \sqrt{\Delta t} \left( \Phi \left( \frac{u_i - \delta}{\sigma \sqrt{\Delta t}} \right) - \Phi \left( \frac{u_i - \delta}{\sigma \sqrt{\Delta t}} \right) \right) \right] + h_{i+1} \left[ (1 - \mu \Delta t) \Phi \left( \frac{u_i}{\sigma \sqrt{\Delta t}} \right) + \mu \Delta t \Phi \left( \frac{u_i - \delta}{\sigma \sqrt{\Delta t}} \right) \right] \right\}, \tag{10}$$

and

$$h_{n-1} = 0. \tag{11}$$

Here, $\Phi$ and $\Phi$ are, respectively, the p.d.f. and c.d.f. of the standard normal distribution. Then, $\{h_i\}$ correspond to the continuation values under the optimal policy.

Suppose further that, for $0 \leq i < n - 1$, $u_i^*$ is a maximizer of (10). Then, a policy that chooses limit order prices that are pegged to the bid prices according to the premia defined by $\{u_i^*\}$, i.e.,

$$l_i^* = S_i + u_i^*, \quad 0 \leq i < n - 1,$

is optimal.

Theorem 1 suggests a computational strategy for determining continuation values and an optimal policy. Starting with the terminal condition $h_{n-1} = 0$, one proceeds via backward induction, solving the single variable optimization problem (10) over the decision variable $u_i$ once per time slot. So long as optimal solutions exist, they will determine the continuation values and optimal policy. Moreover, the optimal policy is a pegging strategy. That is, the limit order price is pegged at a deterministic (but time varying) premium above the current bid price. These limit order premia are given by the maximizers $\{u_i^*\}$.

In the following theorem, whose proof is provided in Appendix B of the electronic companion, we establish the existence and uniqueness of the optimal solutions to (10) and provide upper and lower bounds for the resulting limit price premia, for small values of latency $\Delta t$.

**Theorem 2.** Fix $\alpha > 1$. If $\Delta t$ is sufficiently small, then there exists a unique optimal solution $\{h_i\}$ to the dynamic programming equations (10)–(11). Moreover, the corresponding optimal policy $\{u_i^*\}$ is unique. For $0 \leq i < n - 1$, this strategy chooses limit prices in the range

$$l_i^* \in \left( S_i + \delta - \sigma \sqrt{\Delta t \log \frac{\alpha L}{\Delta t}} , S_i + \delta - \sigma \sqrt{\Delta t \log \frac{R(\Delta t)}{\Delta t}} \right),$$

where

$$L = \frac{\delta^2}{2 \pi \sigma}, \quad R(\Delta t) = \frac{\delta^2 (1 - \mu \Delta t)^2}{2 \pi \sigma^2}.$$
Figure 4. An illustration of the optimal policy of Theorem 2.

Notes. In the absence of latency, at time $t = 0$, the trader would set the limit price at a premium of $\delta$, i.e., $l_0 = S_0 + \delta$. In an environment with latency, the trader might set the limit price to be $l_0^*$, which lowers $l_0$ by an additional safety margin of $C$ standard deviations. This serves to increase the likelihood of trade execution in the interval $(\Delta t, 2\Delta t)$. The optimal limit price $l_0^*$ utilizes a safety margin that is slightly larger.

by an additional safety margin. If we set this safety margin to be $C$ standard deviations of the one-period price change, i.e., $l_0 = S_0 + \delta - C\sigma\sqrt{\Delta t}$, then the probability of missing execution becomes

$$P(l'_0 \geq S_{\Delta t} + \delta) = P(S_0 - C\sigma\sqrt{\Delta t} \geq S_{\Delta t}) = \Phi(-C).$$

This probability can be made close to 0 by the choice of $C$. However, given a fixed choice of $C$ independent of $\Delta t$, the probability remains constant (i.e., independent of $\Delta t$) and nonzero. The additional safety margin corresponding to the log term in Theorem 2 is a second-order adjustment. This is introduced so that, given the optimal limit price $l_0^*$, the probability of execution tends to 1 as $\Delta t \to 0$.

4.2. Asymptotic Analysis

The dynamic programming decomposition developed in §4.1 allows the exact numerical computation of the value $h_0(\Delta t)$, the value under an optimal policy of the latency model introduced in §3, when the latency is $\Delta t$. As discussed earlier, the latency observed in modern electronic markets is extremely small, often on the time scale of milliseconds. Thus, we are most interested in the qualitative behavior of $h_0(\Delta t)$ in the asymptotic regime where $\Delta t \to 0$. The main result of this section is the following theorem, whose proof is provided in Appendix C of the electronic companion. It provides a closed-form expression for $h_0(\Delta t)$, which holds asymptotically as $\Delta t \to 0$.

**Theorem 3.** As $\Delta t \to 0$,

$$h_0(\Delta t) = \tilde{h}_0 \left(1 - \frac{\sigma}{\delta} \sqrt{\Delta t \log \frac{\delta^2}{2\pi\sigma^2\Delta t}}\right) + o(\sqrt{\Delta t}),$$

where

$$\tilde{h}_0 = \delta(1 - e^{-\mu T})$$

is the optimal value for the stylized model without latency, i.e., the value defined by (5).

Theorem 3 is not surprising when considered in the context of Theorem 2. In the stylized model without latency, the optimal strategy is to peg the limit order price at a premium of $\delta$, and this yields a value of $\tilde{h}_0$. On the other hand, Theorem 2 suggests a trader facing latency $\Delta t$ will lower this limit price premium by a factor of, approximately,

$$\frac{\sigma}{\delta} \sqrt{\Delta t \log \frac{\delta^2}{2\pi\sigma^2\Delta t}} + o(\sqrt{\Delta t}).$$

If this lowers the ultimate value proportionally, then the value of the optimal policy in the presence of latency $\Delta t$ should approximately be

$$\tilde{h}_0 \left(1 - \frac{\sigma}{\delta} \sqrt{\Delta t \log \frac{\delta^2}{2\pi\sigma^2\Delta t}}\right) + o(\sqrt{\Delta t}).$$

The proof of Theorem 3, provided in Appendix C of the electronic companion, makes this intuition precise.

One implication of Theorem 3 is that $\tilde{h}_0(\Delta t) \to \tilde{h}_0$ as $\Delta t \to 0$, i.e., the value of the latency model converges to that of the stylized model without latency of §2. This suggests the following definition.

**Definition 1.** Define the latency cost associated with latency $\Delta t$ by

$$\text{LC}(\Delta t) \triangleq \frac{\tilde{h}_0(\Delta t)}{h_0}.$$ (12)

Latency cost has an easy interpretation. Using $\tilde{h}_0$, the value obtained in the stylized model without latency as a benchmark, the numerator of (12) is the lost revenue incurred due the presence of latency. On the other hand, we can regard the denominator as the cost-of-immediacy for an impatient investor in a time horizon of length $T$. This is because, in the stylized model without latency, it is the difference in revenue obtained by a risk-neutral investor willing to patiently provide liquidity by employing limit orders over the length of the time horizon, and an impatient investor who demands immediate liquidity and sells at the bid price at time $t = 0$, cf. (5). Therefore, we can describe the latency cost as the amount a trader forgoes due to latency, as a percentage of the cost-of-immediacy.

The following corollary restates the asymptotic approximation of Theorem 3 in terms of latency cost.

**Corollary 1.** As $\Delta t \to 0$,

$$\text{LC}(\Delta t) \approx \frac{\sigma}{\delta} \sqrt{\Delta t \log \frac{\delta^2}{2\pi\sigma^2\Delta t}} + o(\sqrt{\Delta t}).$$

There are a number of interesting observations that can be made regarding the asymptotic approximation of Corollary 1. First of all, asymptotically, latency cost does not depend on the length of the time horizon $T$ or the arrival rate of impatient traders $\mu$. As a function of the remaining parameters, the asymptotic latency cost depends only...
on a composite parameter that is the ratio the one-period standard deviation of price changes $\sigma/\sqrt{\Delta t}$ to the bid-offer spread $\delta$. Both of these quantities are readily measurable empirically. Corollary 1 suggests that the latency cost increasing in this ratio. Thus, at the same level of latency, the latency cost is most significant for assets that are very volatile or very liquid. Furthermore, Corollary 1 suggests that when latency is low, there are increasing marginal benefits to further reductions in latency, i.e., $\text{LC}'(\Delta t) < 0$. In §5.1, we illustrate some of facts numerically, as well as considering the accuracy of our approximation, as compared to the exact latency cost.

4.3. Discreteness of Time vs. Latency

The latency model introduced in §3 differs from the the stylized model without latency of §2 in two principal ways: (i) the trader faces a delay or latency between the time that trading decisions are made and when they reach the marketplace, and (ii) the latency model is formulated in discrete-time rather than continuous time. The latter point refers to the facts that, in the model with latency, a trader is only able to update his limit order at discrete intervals of time rather than continuously, impatient buyers arrive according to a Bernoulli process rather than a Poisson process, etc. To disentangle these two effects, in this section we will briefly describe a trading model that is formulated in discrete time but without latency. By considering this model, we will demonstrate that the asymptotic latency cost derived in §4.2 is indeed due to latency effects and not due to the discreteness of time.

To this end, consider a model in the discrete-time setting of §3 but with no latency. Here, at each time $T_i = i\Delta t$, for $i = 0, 1, \ldots, n$, the investor sets a limit order price $l_i$. This limit order price takes effect immediately. Between time $T_i$ and time $T_{i+1}$ the following events can occur:

- If $S_{T_i} \leq l_i$, i.e., the bid price is less than the limit order price, the limit order immediately executes at the price $S_{T_i}$.
- Otherwise, suppose that an impatient buyer arrives in the time interval $(T_i, T_{i+1})$ and $l_i \leq S_{T_i} + \delta$, i.e., the limit price $l_i$ is within a margin $\delta$ of the bid price at the start of the interval. In this case, the limit order executes at the price $l_i$. We assume that an impatient buyer arrives with probability $\mu\Delta t$, independent of everything else.

As before, if the investor is unable to sell the share by the end of the time interval, he is forced to sell via a market order, i.e., $l_n = -\infty$. If $P$ is the sale price, the optimal value for the trader in this discrete model is given by

$$h_0^D(\Delta t) \triangleq \max_{l_0, \ldots, l_n} \mathbb{E}[P] - S_0.$$ 

We have the following result, whose proof is identical to the martingale argument used to establish Lemma 1.

**Lemma 2.** An optimal strategy for the discrete model is to place limit orders at the price $l_i = S_{T_i} + \delta$, for $i = 0, 1, \ldots, n - 1$. This strategy achieves the value $h_0^D(\Delta t) \triangleq \delta(1 - (1 - \mu\Delta t)^n)$.

Now, note that for all $0 < \Delta t < 1/\mu$,

$$e^{-\mu T - (1/2)\mu^2 T\Delta t} \leq (1 - \mu\Delta t)^{T/\Delta t} \leq e^{-\mu T}.$$ 

Therefore, the difference in value between the continuous model of §2 and the discrete model considered here is at most

$$|h_0^D(\Delta t) - \bar{h}_0| \leq \delta e^{-\mu T}(1 - e^{-(1/2)\mu^2 T\Delta t}) \leq \frac{1}{2}\delta \mu^2 T e^{-\mu T} \Delta t.$$ 

In other words, this difference is asymptotically $O(\Delta t)$. By Theorem 3, however, the difference between the continuous model and the latency model is asymptotically $\Theta(\sqrt{\Delta t} \log(1/\Delta t))$.

Hence, the asymptotic effect of latency dominates the asymptotic effect of the discreteness of time.

4.4. Extensions

The analysis of the latency model that we have presented proceeded according to two high-level steps:

(i) First, in §4.1, a simplified dynamic programming decomposition was developed. In this decomposition, at each time, the trader’s value function is parameterized by a single scalar, rather than being an arbitrary function of state. This allows the Bellman equation to be solved through a system of $n$ equations in $n$ unknowns, given by (10)–(11).

(ii) Second, in §4.2, an asymptotic analysis of the simplified dynamic programming equations (10)–(11) was performed. This gave rise to the asymptotic latency cost expression of Corollary 1.

The dynamic programming decomposition step (i) that is at the heart of our analysis can be extended to a much broader set of stochastic primitives than the present setting. In each of these cases, a different set of simplified dynamic programming equations, analogous to (10)–(11) would arise, and would require a customized variation of asymptotic analysis step (ii). In particular, consider the following tractable generalizations:

- **Price process.** In our model, the price process $S_t$ is a Brownian motion. Our dynamic programming decomposition only requires that the $S_t$ be a Markov process and a martingale. It would be straightforward to extend the dynamic programming step (i) and consider other Markovian martingales, for example, allowing for non-Gaussian processes, time-inhomogeneous volatility, or for jump processes.

On the other hand, the asymptotic analysis step (ii) we have presented is quite sensitive to distributional assumptions of the price process and would require specialized analysis for any such generalization. In Appendix D of the electronic companion, we consider one generalization of particular interest, where the price dynamics also contain a jump component.
Limit order execution. In our model, the execution of a limit order in the time slot \((T_i, T_{i+1})\) required that the limit order price \(l_{i-1}\) be within a spread \(\delta\) of the bid price \(S_{T_i}\), and that an impatient trader arrive. More generally, our dynamic programming decomposition requires only that the execution of this limit order, conditional on the price difference \(l_{i-1} - S_{T_i}\), be independent of everything else. This can accommodate a number of generalizations, for example, the arrival rate of impatient buyers can be time-varying. Furthermore, the maximum premium above the bid price \(S_i\) that an impatient buyer is willing to pay can be randomly distributed, as in (3). This would allow models where a limit order that is priced aggressively low has a much higher probability of execution. Such models could alternatively be interpreted, as discussed in §2, as cases where the prevailing bid-offer spread is not constant but is independent and identically distributed, varying from period to period.

5. Empirical Estimation of Latency Cost

In this section, we will consider empirical applications of our model. First, we will illustrate the optimal trading policy and the corresponding value function when the model parameters are estimated from high-frequency market data for a single stock. We will also compare the exact latency cost (numerically computed via dynamic programming) to the approximation provided by Corollary 1 in order to access the quality of our approximation. Subsequently, we show the historical evolution of latency cost and implied latency across a range of U.S. equities using cross-sectional data on volatilities and bid-offer spreads during the 1995–2005 period.

Our empirical analysis should be regarded as a first-order study to obtain a rough calibration of our model. It will allow us to analyze the model in relevant parameter regimes, as well as gain a broad understanding of the implications of our model for the trading of U.S. equities. Under our modeling assumptions (e.g., Brownian motion price processes, Poisson arrivals of impatient traders, constant bid-offer spread, etc.), our empirical measurement of latency cost requires estimates of the high-frequency price volatility \(\sigma\) and the prevailing bid-offer spread \(\delta\). Here, we make a number of simplifications and rely on the recent empirical work of Aït-Sahalia and Yu (2009) to obtain these quantities:

- We estimate price volatility \(\sigma\) using the maximum likelihood estimates of the volatility of returns provided by Aït-Sahalia and Yu (2009). Note that this estimation of high-frequency volatility aims to filter out the impact of microstructure noise and obtain an unbiased estimate of daily volatility. However, for an investor with a trading horizon of 1 second, microstructure noise needs to be incorporated as well. Therefore, the high-frequency volatility estimate that is used in our empirical analysis underestimates the actual volatility faced by a high-frequency trader with a very short trading horizon.

- Recall that the prevailing bid-offer spread, \(\delta\), equals the bid-offer spread in the absence of the liquidating trader. In the empirical data, it is impossible to disentangle the presence of liquidating traders. Moreover, the bid-offer spread will not be constant but will vary over the course of the trading day. As a proxy for \(\delta\), we use the average bid-offer spread over the trading day.

Despite these shortcomings, we believe that our empirical analysis can shed light on the importance of latency in the trading of U.S. equities.

5.1. The Optimal Policy and the Approximation Quality

In what follows, we will numerically evaluate the optimal policy in our model, the corresponding value function, and the latency cost approximation. These numerical experiments are meant to be illustrative of our model. We will use realistic model parameters estimated from recent market data for a single stock. Our methodology here is not meant to be authoritative—there are many subtleties in the analysis of high-frequency data; these are beyond the scope of the work at hand. However, we do seek to demonstrate that our model parameters can be readily derived from commonly available data.

Specifically, the model parameters herein are estimated from trade-and-quote (TAQ) data for a stock that is a representative example of a liquid name, Goldman Sachs Group, Inc. (NYSE: GS), on the trading day of January 4, 2010. These data were obtained from the Wharton Research and Data Services (WRDS) consolidated TAQ database. Only trades and quotes originating from the primary exchange (NYSE) during regular trading hours were considered. The model parameters were estimated as follows:

- Initial bid price: \(S_{0}^{GS} = \$170.00\). This was chosen to be the first transaction price on the trading day.
- Bid-offer spread: \(\delta_{GS} = 0.058\), i.e., equivalently, 3.4 basis points relative to the initial price \(S_{0}^{GS}\). This was estimated by computing the average spread between bid and offer quotes over the course of the trading day and rounding to the nearest cent.
- Arrival rate of market orders: \(\mu_{GS} = 12.03\) (per minute). This was estimated by dividing the total number of NYSE trades by the length of the trading day.
- Price volatility: \(\sigma_{GS} = 1.92\) (daily), i.e., approximately equivalent to an annualized volatility of returns of 17.9%. These were estimated from the time series of transaction prices over the course of the trading day, using maximum likelihood estimation as described in Aït-Sahalia and Yu (2009).
- Trading horizon: \(T = 10\) (seconds).

Figure 5 illustrates the optimal limit order policy for GS under different values of latency. If there is no latency, the limit orders are submitted at a constant premium of \(\delta\). When there is latency, the optimal order policy is obtained using the exact dynamic programming solution.
Figure 5. An illustration of the optimal strategy for GS, expressed in terms of limit price premium over the course of time, for different choices of latency.

\[
\sigma \sqrt{\Delta t \log \frac{\delta^2}{2\pi \sigma^2 \Delta t}}.
\]

(13)

In Figure 5, we see that with a latency of 500 ms, this adjustment is up to approximately 1.4 \( \sigma \sqrt{\Delta t} \), i.e., 1.4 times the standard deviation of prices over the latency interval. When the latency is reduced to 250 ms and to 50 ms, the adjustment increases to 1.6 and 2.1 standard deviations, respectively. The fact that this adjustment, when measured as a multiple of the uncertainty over the latency period, increases as the latency decreases is consistent with (13).

In Figure 5, we also observe that as \( t \) increases and the trading deadline approaches, the limit order premium \( u^*_t \) becomes lower. This makes intuitive sense: the trader faced with a terminal value of 0 since he is required to sell using market order at the end of the period. As the deadline approaches, the trader is more willing to sacrifice the potential profits of a limit order in order to increase the probability of execution.

Figure 6 illustrates the corresponding continuation value under the optimal policy for GS, for different values of latency. Clearly, the trader’s expected payoff decreases as latency increases or the end of the trading horizon approaches.

Finally, Figure 7 illustrates the latency cost as a function of latency. Both the exact value of the latency cost, computed numerically via the dynamic programming decomposition (10)–(11), and the asymptotic latency cost approximation provided by Corollary 1 are shown. The latency costs decrease from approximately 20% of the...
cost of immediacy to 5% of the cost of immediacy, as the latency decreases from 500 ms to 5 ms. Furthermore, the marginal benefit of reducing latency increases as the latency approaches zero. Finally, we note that the approximate and exact latency costs are quite close across the entire range of latency values. This suggests that the approximation is of very high quality in this case.

5.2. Historical Evolution of Latency Cost

In this section we will examine the historical evolution of latency cost in U.S. equities. Here, we consider the situation of a hypothetical investor with a fixed latency of 500 milliseconds. This choice of latency is made approximately to reflect the reaction time of a very fast human trader. We will use this as a proxy for the fastest possible trading on a “human time scale.” By analyzing the evolution of the associated latency cost, we will get a sense of the importance of latency over time.

Our empirical analysis relies on the data set of Aït-Sahalia and Yu (2009). Their data set contains estimates for various liquidity measures for all NYSE common stocks on a daily basis during the sample period of June 1, 1995 to December 31, 2005. The estimates are derived from intraday transaction prices and quotes from the NYSE TAQ database. We utilized only the volatility and bid-offer spread data, as we have seen both analytically (Corollary 1) and numerically (Figure 7) that under our modeling assumptions, latency cost can be approximated accurately for low values of latency using only these two measures.

The data set contains volatility and bid-offer spread estimates for given stock on a particular day if the number of transactions on that day exceeds 200. The minimum, average, and maximum number of stocks in the sample on any day are 61, 653, and 1,278, respectively. In particular, earlier periods in the data set contain fewer stocks due to a smaller number of firms and a lower volume of transactions. In this data set, the bid-offer spread is estimated using only NYSE quotes in the regular trading hours. The volatility estimate is obtained using maximum likelihood estimation in the presence of market microstructure noise. Maximum likelihood estimation is preferred over other nonparametric estimation methods (e.g., two scales realized volatility) as a simulation study shows that maximum likelihood estimation provides robust estimators under reasonable stochastic volatility and jump models in the underlying asset. The reader is urged to consult §2.1 of Aït-Sahalia and Yu (2009) for full details of their estimation procedure.

For each stock in the data set, on a daily basis we compute the latency cost facing an investor with a fixed latency of 500 ms using the asymptotic approximation of Corollary 1. These daily latency costs are then averaged over each month. Figure 8 displays percentiles of the monthly averages of latency cost over all the stocks in the sample, as a function of time. As a representative example of a liquid name, we also report the monthly averages of latency cost of Goldman Sachs Group, Inc. (NYSE: GS). Note that the time series for GS begins from its initial public offering in 1999. For reference, we have added an additional point to this time series based on our estimation in §5.1 of the latency cost for GS on January 4, 2010.

Figure 8 illustrates that latency costs have had an increasing trend over the 1995–2005 period. In particular, we observe that the median latency cost incurred by trading on a human time scale roughly tripled, increasing from approximately 5% to approximately 14%. One important factor in this increase has been the reduction of bid-offer spreads over this time period. Instances during the period when the NYSE reduced the tick size (from $1/8 to $1/16 in June 1997, and from $1/16 to $0.01 in January 2001) coincide with spikes in latency cost. This is consistent with bid-offer spreads decreasing significantly and volatility maintaining the same level at these times. This suggests that any future reduction in tick sizes will result in increased latency costs.

Using a data set in a similar time frame, from February 2001 to December 2005, Hendershott et al. (2011) conclude that in the post-decimalization era, the increase in algorithmic trading activity had a positive impact on the level of liquidity. This result suggests that the increase in algorithmic trading in and of itself elevated the importance of low latency trading and increased the cost of latency.

5.3. Historical Evolution of Implied Latency

An alternative perspective on the historical importance of latency comes from considering a hypothetical investor with a target level for the cost of latency, relative to the overall cost-of-immediacy. The representative trader maintains this target over time through continual technological upgrades to lower levels of latency. We determine the requisite level of latency for such a trader, over time and across the aggregate market. In other words, fixing the latency cost percentage LC to the target level, we can solve the asymptotic approximation (12) for the level of latency required at each time to achieve latency cost LC. We call this the implied latency.

Figure 9 illustrates the implied latency values over the 1995–2005 period, assuming that the target level LC = 10% of overall transaction costs result from latency. We observe that the median implied latency decreased by approximately two orders of magnitude over this time frame. The 90th percentile of U.S. equities, for example, went from an implied latency on the scale of seconds to an implied latency on the scale of tens of milliseconds.

5.4. Empirical Importance of Latency

Our model captures the cost of latency due to a lack of contemporaneous information. Figure 8 suggests that, when our model is calibrated to the topmost quartile of U.S. equities, a investor with latency on the human time scale faces a latency cost of at 15% to 25%. To assess the significance of this, we can compare it to other trading costs. Suppose
we normalize the cost of immediacy to $0.01, which is the typical bid-offer spread for a liquid U.S. equity. Then, our model suggests that the benefit of reducing latency from a human time scale of 500 ms to an ultra-low latency time scale of less than 1 ms is approximately $0.0015–$0.0025 per share traded.

While this might seem very small as an absolute number, note that this is also the order of magnitude as other trading costs faced by the most cost efficient institutional investors. For example, a hedge fund would pay an average commission of $0.0007 per share for market access. Furthermore, investors might pay an SEC fee of $0.0005 per share traded, and exchange fees or rebates of $0.0020–$0.0030 per share traded. To the extent that a sophisticated institutional investor is cost sensitive and wishes to optimize these other execution costs, they should also be concerned with latency. This isn’t to suggest that latency cost is important to all investors. A typical retail investor, for example, might pay a brokerage fee that is up to $0.10 per share traded. For this latter type of investor, the cost of latency as described here is not a significant component of overall trading costs.

Alternatively, we can compare the $0.0015–$0.0025 per share traded latency cost to the rents extracted by agents that have made the required technological investments to trade on an ultra-low latency time scale. For example, providers of automated algorithmic trade execution services charge an average commission of $0.0033 per share traded for their execution services, which leverage sophisticated low latency technological infrastructure. Note that this cost is comparable to the latency cost. Another class of agents with ultra-low latency trading capabilities are high-frequency traders. Reported net profit numbers for high-frequency traders are in the range of $0.0010–$0.0020 per share traded. This is of the same order of magnitude as the latency cost.
Figure 9. An illustration of the historical evolution of implied latency over the 1995–2005 time period.

Notes. We consider a hypothetical investor who makes sufficient technological investments to ensure a constant latency cost of 10%. The implied latency is the level of latency required to achieve this latency cost. Percentiles for the implied latency are reported across NYSE common stocks. The implied latencies are computed from the data set of Aït-Sahalia and Yu (2009). The implied latency for GS is also reported, beginning from its IPO. We observe that implied latency has had a decreasing trend over the 1995–2005 period. Specifically, the median implied latency decreased by approximately two orders of magnitude over this time frame.

6. Conclusion and Future Directions

This paper provides a model to quantify the cost of latency on transaction costs. We consider a stylized execution problem, where a trader must sell an atomic unit of stock over a fixed time horizon. We consider this model in the absence of latency as a benchmark, and we incorporate latency by not allowing the trader to continuously participate in the market. Orders submitted by the trader reach the market with a fixed latency, and the trader is forced to deviate from the benchmark policy in order to take into account the uncertainty introduced by this delay. We quantify the cost of latency as the normalized difference in expected payoffs between this model and the stylized model without latency.

Since the latency values observed in modern electronic markets are on the order of milliseconds, we provide an asymptotic analysis for the low latency regime, in which we obtain an explicit closed-form solution. To compute this asymptotic latency cost empirically, we need only to estimate the volatility and the average bid-offer spread of the stock. This is an elegant and practical result because data sets and estimation procedures for these quantities are readily abundant in the literature. Indeed, using an existing data set, we show that the cost of latency incurred by trading on a human time scale (500 ms) increased threefold over the 1995–2005 time frame. In addition, using the alternative approach of keeping a fixed level of latency cost through continuous technological improvements, we compute the various percentiles of the implied latency over this time frame. Using the same data set, we observe that the median implied latency decreased by approximately two orders of magnitude.

Our empirical analysis can also be utilized to compare the magnitude of latency cost to other trading costs...
incurred by institutional investors. Our results suggest that the difference in payoff between trading with a human time scale (500 ms) and an automated trading platform with ultra-low latency (1 ms) is approximately of the same order of magnitude as other trading costs faced by institutional investors. This observation certainly underlines the significance of latency for such investors. In conclusion, our model is the first theoretical approach in the literature to concretely quantify the impact of latency on the optimal order submission policy and its resulting cost to the trader.

There are a number of interesting future directions for research. First, as discussed in §4.4, there are a number of tractable extensions to the present model that can be analyzed. More generally, in the introduction we identified a number of broad themes to the costs that arise from latency. The model we have presented captures mainly costs due to a lack of contemporaneous decision making. It does not capture the latency costs due to strategic effects (i.e., comparative advantage/disadvantage relative to other investors) or due to time priority rules. These remain important questions for future research.

Supplemental Material
Supplemental material to this paper is available at http://dx.doi.org/10.1287/opre.2013.1165.

Endnotes
1. NYSE, pre-1980 upgrade (Easley et al. 2013).
6. Note that the trade quantity of a single share is meant to represent an atomic unit of the asset, or the smallest commonly traded lot size. The underlying assumption is that the desired trade execution will ultimately be accomplished by a single transaction. In typical U.S. equity markets, for example, this atomic unit might be a block of 100 shares.
7. For example, see Bertsimas and Lo (1998) or Almgren and Chriss (2001). These questions have also recently been addressed by Back and Baruch (2007) and Pagnotta (2010) in equilibrium settings.
8. We denote by 1Æ the indicator function of the event Æ.
9. Note that many modern exchanges explicitly allow for pegged orders; these orders obviate the need for the trader to continually track the bid price in the manner we describe. However, more generally, when tracking an alternative informational process such as the price on a different exchange, the fundamental value (see §2), etc. a trader would still need to continuously monitor the market relative to the informational process, and latency would be important.
10. Equivalently, we can assume that our definition of time corresponds to the trader’s clock.
11. In practice, this ordering scheme might be achieved by a sequence of cancel-and-replace limit orders, each of which cancels the prior limit order and inserts a new limit order with the updated price. If the prior limit order has already been filled when a subsequent cancel-and-replace order arrives, the new order will fail. Hence, the investor is guaranteed to sell at most one share.
12. Note that this is simply a discrete-time Bernoulli arrival process that is analogous to the the Poisson arrival process of §2.
13. In what follows, given arbitrary functions f and g and a positive function q, we will say that f(Δt) = g(Δt) + O(q(Δt)) if lim supΔt→0 |f(Δt) − g(Δt)|/q(Δt) < ∞, i.e., if the difference between f and g, as Δt → 0, is asymptotically bounded above by some positive multiple of q. Similarly, we will say that f(Δt) = g(Δt) + o(q(Δt)) if lim supΔt→0 |f(Δt) − g(Δt)|/q(Δt) = 0, i.e., if the difference between f and g, as Δt → 0, is asymptotically dominated by every positive multiple of q. Finally, we will say that f(Δt) = g(Δt) + Θ(q(Δt)) if 0 ≤ lim supΔt→0 |f(Δt) − g(Δt)|/q(Δt) ≤ lim infΔt→0 |f(Δt) − g(Δt)|/q(Δt) < ∞, i.e., if the difference between f and g is asymptotically bounded above and below by positive multiples of q.
15. As of January 21, 2011, the SEC fee is a fraction $0.0000192 of the proceeds of an equity sale. If we assume a typical stock price of $50, this is approximately $0.0010 per share sold. Amortizing this cost equally between buys and sells results in $0.0005 per share traded.
16. For example, at the time of writing, the brokerage firm E-TRADE charges $10 per trade. Assuming a typical trade of 100 shares, this cost is $0.10 per share traded.
17. “U.S. Equity Trading: Low Touch Trends,” TABB Group, July 2010. Note that some institutional investors pay significantly larger commissions for trade execution in order to compensate their brokers for trading ideas or research services. The commission we quote here is for “nonidea driven” services that relate purely to trade execution using the algorithms and technological platform of the broker.

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