

Learning by Investing: Evidence from Venture Capital

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Abstract: Uncertainties about technologies and investment opportunities are prevalent for investments in entrepreneurial companies by venture capitalists (VCs), and this study finds that the resolution of these uncertainties, through VCs' learning, is important for their investment decisions. The hypothesis that individual investments are evaluated in isolation, as predicted by standard models, is clearly rejected. The empirical analysis is based on a dynamic learning model derived from the Multi-armed Bandit model. The results suggest that VCs learn from past investments (exploitation) but also consider the option value of future learning (exploration) when making investment decisions.

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Given their importance for financing high-tech entrepreneurial companies, venture capitalists (VCs) have substantial impact on innovation and development of new technologies (Kortum and Lerner (2000)). While a great deal has been written about the relationship between VCs and their portfolio companies,¹ less is known about VCs' decisions to invest in particular industries and companies.² These investments are fraught with uncertainty (Quindlen (2000)), and to understand VCs' investment decisions, it is important to understand how these uncertainties are resolved, i.e. how VCs learn. This study finds that VCs' investment decisions are affected both by the expected return from the investment itself but also from the potential to learn from it. Learning is valuable, since it improves the investors' understanding of the various investment opportunities and improves their future decisions. The hypothesis that VCs' investments are chosen independently to maximize the return from each investment individually, as predicted by standard models, is clearly rejected.

The empirical evidence is found using an empirical model derived from the Multi-armed Bandit model (see Berry and Fristedt (1985) and Gittins (1989)). The central ingredient in this model is the investors' beliefs about the profitability of their investment opportunities, and the dynamics of these beliefs. These beliefs are shaped by their past investments and their outcomes, and they affect investment decisions in two ways. First, not surprisingly, investors prefer investments with greater expected immediate returns.

¹ See, for example, Gorman and Sahlman (1989), Sahlman (1990), Lerner (1995), Gompers and Lerner (1999), Hellmann and Puri (2000), Hellmann and Puri (2002), Kaplan and Strömberg (2004), and Hochberg, Ljungqvist and Lu (2006).

² Notable exceptions are Kaplan and Strömberg (2004) and Gompers, Kovner, Lerner and Scharfstein (2005).

Without learning, the immediate return is the only return, and this is what is usually captured in financial models. With learning, an additional, indirect, effect arises since investors prefer investments with more informative outcomes, because these investments help them learn and improve their future decisions. This value can be viewed as an option value of learning, and it generates a trade-off between *exploiting* investments with known payoffs and *exploring* investments with uncertain payoffs.³ The main contribution is to provide an empirical methodology for separating and measuring exploration and exploitation. To my knowledge, this is the first time these two effects have been estimated separately, and the main results confirm that both of them are indeed important determinants of VCs' investment decisions.

Additional predictions from the model are that more valuable investments are made faster and VCs that explore more are more successful. These predictions are also confirmed empirically, lending further support to the main results. Methodologically, the model relies on a statistical *index result* to simplify the empirical analysis. Previous studies of learning⁴ use computationally intensive estimation procedures to capture the inherent dynamic programming problem (see Crawford and Shum (2005) and Erdem and Keane (1996)). Relying on the index result leads to more transparent and tractable inference, allowing the model to be estimated using standard statistical procedures.

³ The terminology of *exploitation* and *exploration* is introduced by March (1991) in the context of organizational learning.

⁴ Another strand of the learning literature considers learning-by-doing, where learning is a free byproduct of an activity (Arrow (1962)). Recent finance applications include Pastor, Taylor and Veronesi (2006), Linnainmaa (2006), and Hochberg, Ljungqvist and Vissing-Jorgensen (2008), who study learning by the limited partners in VC funds.

Previous applications of the bandit model include Rothschild (1974)'s model of firms' experimentation with prices to learn about uncertain demand and Weitzman (1979)'s analysis of optimal sequencing of research projects. Manso (2006) study incentive provision in a learning model similar to the Bandit model, and Bergemann and Hege (1998) and Bergemann and Hege (2005) present theories of staged financing based on the bandit model in the context of VC investments. Empirically, Jovanovic (1979) and Miller (1984) estimate models of job turnover in which workers learn about job-specific skills.

Taking the Bandit model to the data requires additional assumptions. As a starting point, the model assumes that investors chose between investments at the industry level, and that learning takes place at this level as well. This is a natural starting point, and it is motivated both by data limitations and by other assumptions inherent in the Bandit model. Goldfarb, Kirsch and Miller (2007) find evidence of VC behavior consistent with learning at the industry level. The data limitations arise since inference about learning behavior is derived from the investment histories across VCs. The model links past investments and their outcomes to subsequent investments within the same category. Defining categories at finer levels leads to more categories and shorter investment histories in each category, but the estimation procedure requires fairly long histories to reliable inference.

Further, the model assumes that the environment is stationary, that investors only learn from their own past investments, and that investments in one industry are not informative about investments in other industries. These assumptions are not entirely

reasonable in the context of entrepreneurial investing, but they are probably more palatable at the industry level than at finer levels. The assumptions are technical assumptions that are required for the index result, which, in turn, is required for the tractability of the empirical analysis tractable. The severity of these assumptions is somewhat mitigated by including market-level controls in the empirical analysis. These controls capture general trends in the market and the effect of learning from other investors. However, this is a reduced form solution, and a more formal relaxation of these assumptions requires explicitly solving the investors' dynamic programming problems as part of the estimation procedure, making the inference problem largely intractable. With the assumptions, the index result overcomes this.

After investing in a company, the outcome of each investment is either a success (subsequent IPO or acquisition of the company) or a failure (liquidation). While this is a coarse outcome measure, it is difficult to obtain more detailed information about the financial performance of these investments, and this classification is standard in the literature. Binary outcomes also simplify the updating of investors' beliefs, since these beliefs are about the distribution of the success probability, which is easily captured by the Beta distribution.

The paper proceeds as follows. The following section presents the theoretical learning model and the index result. The second section presents the data and variables, and discusses the econometric implementation of the model. Section three presents the empirical evidence of learning. Section four discusses the construction and interpretation of the investors' prior beliefs, and the final section concludes.

I. The Multi-Armed Bandit Model

In the Multi-armed Bandit model⁵ an investor faces an infinite sequence of periods, $t = 0, 1, \dots$. Each period the investor chooses between K arms, denoted $i = 1, 2, \dots, K$, where each arm represents an investment in an entrepreneurial company in this industry. The outcome of an investment is either a success or a failure, as indicated by $y_i(t) \in \{0,1\}$, and the success probability, given by $p_i = \Pr[y_i(t) = 1]$, is constant over time but can vary across industries and investors. The investor does not know p_i but has prior beliefs, given by $F_i(0)$. These beliefs are updated after each investment, using Bayes rule, and the updated beliefs before investing at time t are $F_i(t)$. The support of $F_i(t)$ is the interval from 0 to 1, representing all possible values of p_i .

With binary outcomes, the assumption that investors' initial beliefs are distributed $Beta(a_0, b_0)$ helps simplify the Bayesian updating. Let r_i be the number of past successes and n_i be the total number of past investments in industry i . The updated beliefs are then simply $Beta(a_i, b_i)$, where $a_i = a_{i,0} + r_i$ and $b_i = b_{i,0} + n_i - r_i$. In other words, a_i counts the number of past successes and b_i counts the number of past failures. As the number of investments increases, the mass of the distribution becomes concentrated at the empirical success rate, defined as $\lambda_i = a_i / (a_i + b_i)$, which also equals the mean of the $Beta(a_i, b_i)$ distribution. In other words, given the beliefs, λ_i is the expected value of p_i , and this

⁵ See Berry and Fristedt (1985), Gittins (1989), and Whittle (1982) for general discussions of this class of models.

value also equals the investors' expected immediate return from the investment, since

$$E[y_i | F_i(t)] = \Pr[y_i = 1 | F_i(t)] = E[p_i | F_i(t)] = \lambda_i.$$

The investor's strategy specifies investments as a function of past investments and their outcomes. The strategy at time t is $s(t) : \{1, \dots, K\}^t \times \{0, 1\}^t \rightarrow \{1, \dots, K\}$, and the full strategy is $S = \{s(0), s(1), s(2), \dots\}$. The investor's problem is to determine the strategy that maximizes total expected return. Let δ denote the discount factor, and the investor then solves

$$V = \sup_S E \left[\sum_{t=0}^{\infty} \delta^t y_{s(t)}(t) \mid F(0) \right]. \quad (1)$$

Formulated as a dynamic programming problem, the Bellman equation is

$$V(F(t)) = \max_{s(t)=1,2,\dots,K} E \left[y_{s(t)}(t) \mid F(t) \right] + \delta E \left[V(F(t+1)) \mid F(t), s(t) \right]. \quad (2)$$

where the state variables contain the investor's updated beliefs. These develop according to the transition rules

$$F_i(t+1) = F_i(t) \text{ for } s(t) \neq i \quad (3)$$

$$F_i(t+1)(v) = \begin{cases} \text{Beta}(a_i + 1, b_i) & \text{for } s(t) = i \text{ and } y_i(t) = 1 \\ \text{Beta}(a_i, b_i + 1) & \text{for } s(t) = i \text{ and } y_i(t) = 0 \end{cases} \quad (4)$$

Equation (3) states that the beliefs are unchanged unless an investment is made in an industry, and equation (4) reflects Bayesian updating of the beliefs about p_i after investing in industry i and observing either $y_i(t) = 1$ (success) or $y_i(t) = 0$ (failure).

A. Gittins Index

The Bandit problem is a difficult dynamic programming problem, due to the high dimensionality of the state space. With six industries and beliefs captured by two variables, the resulting state space is twelve dimensional, which give a numerically challenging dynamic programming problem, in particular when this problem is solved repeatedly as part of an estimation procedure. A breakthrough was made when Gittins and Jones (1974) derived the solution to this problem, formulated in terms of the *Gittins Index*. This index is calculated separately for each industry and summarizes all relevant information, and their *index result* shows that the optimal strategy is to simply choose the industry with the highest value of the index. Let $v_i(t)$ be the index for industry i , at time t , and the optimal strategy is then

$$s(t) = \arg \max_{i=1, \dots, K} v_i(t). \quad (5)$$

The Gittins index is central for the analysis below. While there is no (known) closed form solution to the index, it can be calculated numerically by solving a reduced dynamic programming problem for each industry separately (see Gittins (1989)). To understand the properties of the index, it is helpful to review an approximation, derived by Gittins and Jones (1979), on the form

$$v(a_i, b_i) = \lambda_i + \text{Option Value}(\lambda_i, n_i), \quad (6)$$

where *Option Value* is a non-negative tabulated function, and $\lambda_i = a_i / (a_i + b_i)$ and $n_i = a_i + b_i$. The total present value of an investment is the total value of the index, and

this value can now be broken down into two parts. The first term, λ_i , is the expected immediate return. This is the value of the investment without any learning, and clearly the total value is at least this large. The second term is the value of the investment in excess of the immediate return, and this represents the value of information or the option value of learning. This value is illustrated in Figure 1 for three levels of n_i . Holding n_i fixed, *Option Value* is greater for intermediate values of λ_i and vanishes as λ_i approaches zero or one. For given λ_i , the option value tends to zero as n_i increases. Intuitively, when the number of investments increases, beliefs become more informed and tighter distributed around λ_i , which in turn approaches the true p_i , and the value of learning vanishes.

*** FIGURE 1 HERE ***

B. Interpretation of Learning

The assumption that investors learn about success probabilities across industries captures several kinds of learning. Investors may be uncertain about their private abilities. For example, access to deal flow and the ability to screen and work with entrepreneurs are important skills for VCs. These abilities differ across investors and industries, and new VCs may not fully know their own abilities. Further, they learn about current industry conditions. Historically, VCs investments have seen numerous cycles where, i.e. hardware, web-based software, and biotechnology have gone in and out of fashion, and the investors learn about these cycles from their current investments. Finally, the model captures learning at the finer levels of individual ideas or technologies,

aggregated to the industry level. New technologies and business models are fraught with uncertainty, and learning about the viability of those is closely related to Arrow (1969)'s definition of innovation, stating that “[t]echnological progress is in the first instance the reduction in uncertainty.” The option to make follow-up investments in other companies with similar technologies may be a substantial part of the value of an investment in an entrepreneurial company and a driver of technological progress.

Two problems prevent the model from explicitly capturing learning at the level of individual technologies or ideas. The estimation requires a fairly large number of observations of investments in each category, and defining categories at a finer level reduces these numbers. Further, the assumptions that the environment is stationary and that investments are only informative about other investments in the same category become more problematic at finer levels. However, without these assumptions, the index result fails, severely complicating the analysis.

The model also assumes private learning, and excludes learning by observing other investors or observing the market in general. Again, this assumption is necessary for the index result, but the severity of the assumption is mitigated somewhat by including market-level controls in the empirical specifications. Note that the ability to observe and learn from other investors may give rise to additional informational frictions. A free-rider problem arises when investors can learn from each other, since investors have an incentive to reduce exploration, knowing they can benefit from others' investments in learning (see Bolton and Harris (1999), and Keller, Rady and Cripps (2005)). An informational herding problem arises when investors observe others'

investments and this public information lead them to rely less on their own private learning (see i.e. Morris and Shin (2002) and Amador and Weill (2006)). In both cases, informational frictions reduce investors' incentives to internalize the value of their private information, which reduces exploration and learning below the first-best levels.

II. Empirical Implementation

A. Description of Sample of VC Investments

The data are provided by Sand Hill Econometrics (SHE) and contain the majority of VC investments in the U.S. in the period 1987 to 2005.⁶ SHE combines and extends two commercially available databases, Venture Xpert (formerly Venture Economics) and VentureOne. These are extensively used in the VC literature (i.e. Kaplan and Schoar (2005) and Lerner (1995)), and Gompers and Lerner (1999) and Kaplan, Sensoy and Strömberg (2002) investigate the completeness of Venture Xpert and find that it contains most investments and that missing investments tend to be less significant ones.

The sample is restricted to investments made before 2000, since it typically takes companies three to five years after the initial investment to go public or be acquired, and information about these outcomes is current as of 2005. It is common for multiple VCs to invest in the same company, and the sample contains these multiple investments. VCs

⁶ It may be a concern that only few companies go public after 2000. For robustness, the model is estimated restricting the sample to end in 2000, 1998, 1996, 1994, and 1992. The main results are robust across these sub periods. The signs and economic magnitudes of the main coefficients (unreported) are largely unchanged, and although the statistical significance is reduced with the smaller sample size, the main coefficients remain statistically significant.

also typically stage their investments, but the sample is restricted to each VC's initial investment in a company, to focus the analysis on learning from individual companies and the effect on subsequent investments in other companies. While it would be interesting to study learning from individual rounds, the absence of round level outcome measures prevents this. Further, VCs that make less than 40 investments in the full sample are excluded, since their short investment histories make it difficult to draw inference about their learning process and this create convergence problems for the estimation procedure. This reduces the sample from 3,364 to 216 VCs and eliminates around 50% of the companies. Not surprisingly, eliminated VCs are smaller and more idiosyncratic with lower success rates than the remaining ones. The average success rate for the investors in the final sample is 50% (see Table I below), and the corresponding rate for eliminated investors is only 39%. Overall, the final sample contains 19,166 investments in 6,076 companies by 216 VC firms.

B. Company Characteristics

Each company is classified as belonging to one of six industries. These are "Health / Biotechnology," "Communications / Media," "Computer Hardware / Electronics," "Software," "Consumer / Retail," and "Other,"⁷ and the distribution of investments across industries is presented in Table I. These classifications are aggregated from twenty-five minor classifications. This aggregation is necessary to ensure that

⁷ Since learning may be less pronounced for investments in the "Other" category, the model is also estimated while excluding companies in this category. The empirical results (unreported) are unchanged.

investors have sufficiently long investment histories within each industry classification, and the minor classifications are aggregated with the intent that experience in one industry is informative about subsequent investments in this same industry but not across industries, as assumed in the model.

**** TABLE I ABOUT HERE ****

For each investment, the outcome is given by the binary variable *Success*, and for each investor the variable *Success Rate* measures performance as the number of past successful investments divided by the total number of past investments. An investment is successful when the company eventually goes public or is acquired. This classification is consistent with VCs generating most of their returns from a few successful investments, and is common in the VC literature. Ideally, success would be measured in dollars or as a percentage return, but these data are not widely available. To address concerns about the robustness of this binary outcome measure, Gompers and Lerner (2000) compare different measures, including counting acquisitions as unsuccessful, and they find that these measures are highly correlated and lead to qualitatively similar results. In Table I, panel B, the average success rate is 50.3%, ranging from 13.3 to 86.4% across the VCs in the sample.

For each investment the company is classified as either early-stage or late-stage, where late-stage roughly corresponds to the company having regular revenues. The binary variable *Stage* equals one for late-stage companies, and 28.7% of the investments are in such companies.

C. Market Conditions

In addition to learning from their own investments, investors may be affected by general public signals and market conditions. Two variables are used to control for these effects. *Industry Investments* measures the total number of VC investments in each industry in each year. It varies from 36 investments in “Other” in 1994 to 3,443 investments in “Computer Hardware / Electronics” in 2000. Following Gompers, Kovner, Lerner and Scharfstein (2005), the variable *Industry IPOs* contains the number of VC-backed companies going public in each industry for each year. Industries with more IPOs may represent more profitable investment opportunities and attract more VCs. Gompers, Kovner, Lerner and Scharfstein (2005) find that this is an important determinant of investments, and the results below confirm their finding.

D. Calculating Expected Returns and Option Values

For each investment the investor’s expected immediate return and the option value of learning are calculated from the updated beliefs. Initial beliefs are assumed distributed $Beta(a_{i,0}, b_{i,0})$ with $a_{i,0} = 1$ and $b_{i,0} = 19$, and this particular choice is described in detail below. The updated beliefs are found by counting the number of previous successful and unsuccessful investments, as discussed above, and the resulting counts are given by a_i and b_i , respectively. The expected return is calculated as $\lambda_i = a_i / (a_i + b_i)$. In Table I, this return is found to equal 26.0% on average, varying between 3.2 and 71.7%.

The option value is calculated by subtracting the expected return from the Gittins index. I assume a discount factor of $\delta = 0.99$. With an average time between

investments of 48 days, this corresponds to an annual rate of 8% (the results are robust across a wide range of discount factors). The calculation is performed using a numerical algorithm from Gittins (1989),⁸ and *Option Value* is found to be 6.2% on average, ranging from 1.9 to 8.2%, as reported in Table I. As a fraction of the total value of the investments, the option value varies from 3.1 to 53.1% with an average of 25.2%.

E. Econometric Specification

VCS' investment decision across industries is specified as a multinomial discrete choice model. At the time of each investment, the value of the investment in industry i is given by

$$v_i = \lambda_i \beta_1 + \text{Option Value}_i \beta_2 + X_i' \beta_3 + \varepsilon_i. \quad (7)$$

The investor picks the industry with the highest total value, and the probability that an investor invests in industry i is denoted π_i . When ε_i follows an *i.i.d.* extreme value distribution, the model is equivalent to a Multinomial Logit model (see McFadden (1973) and McFadden (1974)), and it is well known that

$$\pi_i = \frac{\exp(v_i)}{\sum_{k=1, \dots, K} \exp(v_k)}. \quad (8)$$

In this model the scale of the coefficients is not identified, and it is normalized by fixing the variance of the error term to equal one. Note further, that the general quality of the

⁸ A MatLab program for this calculation is available from the author.

investor is not identified either. An investor fixed effect entering across investments cancels out of the maximum operator, and the coefficients are estimated from the relative values of these investments. This means that the model cannot measure an investor fixed effect, but the estimates are consistent with the presence of such an effect, even if its expected value changes over time (i.e. if investors learn about their fixed effect in addition to their industry specific p_i).

To interpret the parameters, a high value of β_1 means that investors are more likely to invest in industries with higher expected immediate returns (higher λ). In other words, β_1 captures investors' tendency to *exploit*, and the Bandit model predicts that this coefficient is positive. A high value of β_2 indicates that investors place weight on the option value and invest in companies with a greater value of information, corresponding to more explorative behavior. Finally, the model predicts that the optimal balancing of these two effects is achieved when $\beta_1 = \beta_2$.

III. Evidence of Learning

Estimates of several specifications of equation (7) are reported in Table II. Across all specifications the coefficients on both λ_i and *Option Value_i* are positive and significant. Not surprisingly, investors *exploit* and prefer industries with a higher expected immediate return (i.e. $\beta_1 > 0$), but investors also *explore* and invest in industries with a greater value of learning (i.e. $\beta_2 > 0$).

**** TABLE II ABOUT HERE ****

Specifications 2 to 5 control for other factors that may affect investment decisions. Although not part of the formal model, investors may learn from other investors and from public market signals. Specifications 2 and 3 include additional controls for the total number of VC-backed IPOs and total number of VC investments during each year in each of the six industries. These market conditions have small but positive and significant effects on investment decisions, and this is consistent with Gompers, Kovner, Lerner and Scharfstein (2005) who document that investors follow public market signals. However, including these additional controls does not eliminate the effects of λ_i and *Option Value_i*, and after controlling for the general trends in the market, investors still internalize the option value of learning.

Specification 4 includes the investors' experience in individual industries (*Industry Experience*), calculated as the total number of past investments the VC has made in each industry. This captures other factors, besides learning, that may lead VCs to concentrate investments in industries where they have longer investment histories and more experience. One can imagine VCs "entrenching" themselves in industries, perhaps to enjoy wider access to the deal flow and excluding entrants, regardless of the specific outcomes of their past investments. Hochberg, Ljungqvist and Lu (2007) find evidence that the syndication practices facilitate this as well. Finally, since industry experience is inversely correlated with option value (a greater experience leaves less scope for learning), this variable may capture misspecifications of the learning process. Still, if these were severe, the multicollinearity arising when including both *Option Value* and *Industry Experience* would reduce the statistical significance of the estimated

coefficients, but the results in Table II show that the coefficients on λ and *Option Value* remain positive and significant when this additional variable is included.

Specification 5 is a kitchen-sink regression with all regressors. The variable *Previous* is a binary variable that equals one for the industry of the investor's previous investment. This captures additional persistence in investment decisions. For example, if investors only partly update their beliefs between investments, they would be more likely to invest in the same industry again, and the coefficient on *Previous* would be positive (conversely, if they spread their investments across industries before updating their beliefs, it would be negative). Overall, the positive and significant coefficients on *Industry Experience* and *Previous* reveal some degree of persistence unexplained by the model. Controlling for these effects, the positive and significant coefficients on *Industry IPOs* and *Industry Investments* show that investors still follow general trends in the market and public market signals. While including these additional determinants of investments, the coefficients on λ and *Option Value* remain positive and significant, confirming that learning is a significant determinant of investment decisions.

To provide further evidence for the model, Table III presents estimates of a Probit model where the outcome of each investment is a function of the investor and market characteristics. Specification 1 is the baseline specification. The results show that investments with higher λ , higher *Option Value*, and in late-stage companies are more likely to be successful. In the learning model, λ measures the investor's expected immediate success probability, and the positive and significant coefficient reported in Table III confirms that this measure captures a reasonable amount of the actual success

probability, indicating that investors' beliefs in the model capture economic aspects of the outcomes of the actual investments. The model also predicts that exploratory investments, which are made mainly for their option value, have lower success rates. An investment that is attractive because $\lambda + \textit{Option Value}$ is large should only have a success rate of λ , and *Option Value* should be a weaker predictor of success. Note that this is not a formal prediction of the model. If investors underestimate the value of untried industries, investments with higher *Option Value* have better realized outcomes. Still, the coefficient on *Option Value* provides some indication of the fit between model and beliefs. Specification 2 includes industry controls and controls for the investor's experience, and the significance of *Option Value* decreases substantially. Finally, specification 3 is a kitchen-sink regression with a number of additional controls for market conditions and investor experience. In this specification the statistical significance of *Option Value* vanishes entirely, but the significance of λ remains largely unchanged, consistent with model.

**** TABLE III ABOUT HERE ****

A. *Investment Strategies and Outcomes*

The model further predicts that investors that explore and learn more have higher success rates. To confirm this relationship the model is estimated separately for each investor. For this estimation, the value of an investment is specified as

$$v_i = [\lambda_i + \textit{OptionValue}_i] + \textit{OptionValue}_i \gamma_{j,1} + \textit{Industry IPOs}_j \gamma_{j,2} + \varepsilon_i. \quad (9)$$

The bracket contains the immediate return plus the option value, i.e. the Gittins index. The second term is option value in excess of the value in the bracket, and investors with positive $\gamma_{j,1}$ exhibit more explorative investment behavior than predicted by the model. Finally, the coefficient $\gamma_{j,2}$ classifies the investment behavior according to how closely it follows general trends in the market, here measured by *Industry IPOs*.⁹ A larger value of $\gamma_{j,2}$ corresponds to an investor that follow these trends to a greater extent.

One slightly unusual feature of this specification is that the scale of the equation is normalized by fixing the “coefficient” for the first term (in the bracket) to equal one.¹⁰ Usually, for discrete choice models, the scale is normalized by fixing the variance of the error term. However, by normalizing the term in the bracket it is possible to estimate the standard deviation of ε_i , and the corresponding coefficient, denoted σ_j , measures how closely the investors follows the prediction of the model, and it can be interpreted as a measure of the investor’s “randomness” or “opportunism.” This normalization also makes the estimated coefficients comparable across investors. With the usual normalization, the randomness is captured by the scale of the coefficients, and an investor making more random investments would have smaller coefficients in absolute terms, making the coefficients difficult to compare across investors.

⁹ The results are largely similar when general trends are measured using *Industry Investments*.

¹⁰ Note that since the scale of this equation is not identified, this restriction is equivalent to the restriction that the two coefficients are just equal (and not necessarily equal to one). The difference is that the scale is now captured by the variance of the error term, not by the magnitude of the coefficients. As explained, this helps make the estimated coefficients comparable across investors.

Technically, the model is estimated by first estimating a standard Logit model and then rescaling the coefficients using the value of the first coefficient. The standard error of the first coefficient provides a measure of how precisely the characteristics of the investor's strategy are estimated, and for investors with shorter investment histories the coefficients are less precise. To adjust, investors are weighted according to the precision of the estimates of their characteristics, with less weight placed on investors with less precise coefficients. One disadvantage of this method is that a small number of investors have negative estimates of the first coefficient of the model, due to random sampling. These are typically investors with short investment histories and imprecisely estimated characteristics. However, these investors appear to have negative values of σ_j , which is difficult to interpret. Since these investors have low weights, all the results are robust to excluding them as well as replacing σ_j with its absolute value. To have interpretable magnitudes of the coefficients, these measures are standardized to have standard deviations equal to one in the sample. Panel B in Table I presents both the scaled and raw estimates.

Treating each investor as an individual observation, and using *Success Rate* as the performance measure, estimates of the following regression are reported in Table IV.

$$Success Rate_j = \beta_0 + \gamma_{j,1}\beta_1 + \gamma_{j,2}\beta_2 + \sigma_j\beta_3 + \varepsilon_j \quad (10)$$

A positive estimate of β_1 indicates that investors that explore more by placing more weight on *Option Value* have higher success rates, and this is the prediction of the model.

A positive estimate of β_2 indicates that investors that follow general trends more have

higher success rates, and the model has no prediction about this parameter. A negative estimate of β_3 indicates that investors that deviate more from the model have lower success rates, and this may be supportive of the model, since the model prescribes the optimal decisions given the beliefs.

**** TABLE IV ABOUT HERE ***

In the first specification in Table IV, panel A, investors with higher $\gamma_{j,1}$ have higher success rates, consistent with the model. An investor that explores more discovers more successful investments and realizes a higher success rate. The greater propensity to explore may be a result of more dispersed prior beliefs or a higher discount factor (a δ closer to one), leading to a higher value of learning. Alternatively, the explorative behavior may be a result of suboptimal investment decisions, but even in this case “excess” exploration should lead to a higher success rate. The coefficient on *Standard Deviation* shows that investors with a higher σ_j have consistently lower success rates, suggesting that investors who deviate more from the learning model or make more “random” or “opportunistic” investments are less successful. The magnitudes of the effects of *Option Value* and *Standard Deviation* are economically meaningful. A one standard deviation increase in exploration (within the sample of investors) is associated with a 2.14 to 2.62% increase in success rate, and a one standard deviation in the “randomness” is associated with a 1.60 to 2.40% drop in success rate. Compared to an average success rate of 50.3% in the sample, these are meaningful effects.

The second specification includes $\gamma_{j,2}$, which has a positive but insignificant coefficient, suggesting that the relationship between investors' performance and their tendency to follow the market is weaker. Specification 3 also includes the total number of investments by the investor (*Final Experience*),¹¹ and the estimated coefficient is positive but insignificant. Kaplan and Schoar (2005) and Sørensen (2007) find that more experienced VCs make more successful investments, but the evidence here is less conclusive, perhaps due to the right-truncation of the sample, making *Final Experience* a noisy measure of experience.

The outcomes of the individual investments provide a finer view of the learning process. Treating each investment as a separate observation, Panel B in Table IV reports the coefficients for a Probit model where the outcome of each investment is estimated as a function of the investor's strategy and additional controls. The empirical specification is

$$\Pr(\text{Success}_i = 1) = \Phi\left(\beta_0 + \gamma_{1,j}\beta_1 + \gamma_{2,j}\beta_2 + \sigma_j\beta_3 + X'_{i,j,t}\beta_4\right). \quad (11)$$

The results confirm that investments by more explorative investors are more successful, and investments by more "random" or "opportunistic" investors are less successful. This is consistent with the evidence from the investors' success rates, and the economic effects are similar. A one standard deviation increase in $\gamma_{1,j}$ corresponds to an

¹¹ The difference between *Final Experience* and *Total Experience* is that *Final Experience* is calculated once for each investor at the end of the sample. *Total Experience* is calculated at the time of each investment and increases through the sample period.

increase in the success probability from 1.48 to 3.35%, and a similar increase in σ_j corresponds to a decrease in the success probability of 2.12 to 2.99%. The second specification includes the measure of the investor's tendency to follow the market, but this effect is again small and insignificant. The final specification is a kitchen-sink regression with additional controls and fixed effects. Not surprisingly, investments in companies at the late stage are more likely to be successful (15.14%). Investments by more experienced investors are marginally more likely to be successful and investments in industries with more VC-backed IPOs are marginally less successful, which is again consistent with Kaplan and Schoar (2005). Overall, the results at the investment level supports the evidence at the investor level, although the sign on $\gamma_{2,j}$ reverses in the last specification.

B. Investment Speed

The model also has implications for the timing of the investments. There are several possible hypotheses. Investors may initially make slow explorative investments, and if these investments are successful, accelerate to benefit from their informational advantage. Alternatively, investors may make quick initial explorations, perhaps to capture first-mover advantages, and then continue at a more measured pace. While the model does not explicitly incorporate investment speed, it provides a framework to investigate this prediction. In the model, speed is determined by the discount factor. A δ closer to one implies less discounting between investments, equivalent to a greater speed. As a starting point, assume that increasing the speed requires costly effort. This

may reflect the cost of searching for new investments or investing in lower quality companies and working harder to improve them. The investor's problem is

$$V(F(t)) = \max_{s(t), e} \left[E y_{s(t)}(t) | F(t) \right] - C(e) + \delta(e) E \left[V(F(t+1)) | F(t), s(t) \right]. \quad (12)$$

Here e is effort, $C(e)$ is an increasing convex cost of faster investing, and $\delta(e)$ is the discount rate, which tends to one as effort increases. This extended model predicts that when the continuation value increases (the last term in equation (12)), the benefit of speed also increases, regardless of whether the continuation value reflects a greater value of learning or a larger immediate return.

This is confirmed empirically. The coefficients of the following OLS regression are reported in Table V.

$$Time_i = \beta_0 + Gittins_i \beta_1 + X_i' \beta_2 + \varepsilon_i. \quad (13)$$

Time is the number of days since the investor's previous investment, and a longer time is equivalent to a slower speed. To control for investments that are made simultaneously, the sample is restricted to investments that are made at least fourteen days apart.¹² In this sample, the average of *Time* is 80.8 days with a standard deviation of 126. The variable *Gittins* represents the continuation value, given by the investment's Gittins index.¹³

¹² The regression results are similar when all the investments are included, but the hazard model described below has problems estimating the hazard rates for investments that are very close.

¹³ Formally, the continuation value is the Gittins index scaled by a factor, see Whittle (1982) (p. 214). Notice that a formal solution to this problem would adjust the continuation value to capture the expected future speed and its cost. This problem is not solved here.

In the first specification in Table V the coefficient on *Gittins* is -119.85. As predicted, investments with greater values are made faster. Specification 2 includes industry and year controls, along with the investor's total experience, and more experienced investors are found to invest faster, although the magnitude of this effect is smaller.

**** TABLE V ABOUT HERE ***

In specifications 3 and 4 the option value and the immediate return enter separately. The model predicts that both coefficients should be negative with similar coefficients, but in specification 3 the sign of the coefficient on *Option Value* is positive. However, in specification 4, when including year and industry controls, both coefficients become negative, although their magnitudes are fairly different. The larger negative coefficient on *Option Value* suggests that investors make explorative investments faster, perhaps to capture first-mover advantages or for other reasons outside the model.

Finally, the investment speed can be captured by a hazard model. In Table V, specifications 5 and 6 report estimates of a Cox hazard model and the results are consistent with the results from the previous specifications. Note, for the hazard model, coefficients greater than one reflects an increase in the hazard rate, corresponding to a shorter time between the investments (and corresponding to a negative coefficient in the OLS regressions). Again, more valuable investments are made quicker, and this effect is observed for both investments with higher immediate returns and higher option values of learning.

Overall, the supporting evidence confirms the predictions of the model. For the investment speed, it is remarkable that the expected returns (λ) and option values are calculated independently of the timing of the investments. The relationship between these variables lends strong support for the model, and it is difficult to find alternative explanations for these findings combined. The variables *Option Value* and λ do indeed appear to measure meaningful aspects of the investors' beliefs, as reflected in both their investment speed and industry choices.

IV. Specification and Interpretation of Prior Beliefs

A. Specification and Interpretation of Prior Beliefs

For the estimation it is assumed that investors' prior beliefs are distributed $Beta(a_0, b_0)$ with $a_0 = 1$ and $b_0 = 19$. These beliefs are determined from the restriction in the model that investors place equal weight on immediate returns and option value, and hence $\beta_1 = \beta_2$. To impose this assumption, a grid search is performed over possible values of a_0 and b_0 . For each value the model in equation (7) is estimated and the restriction $\beta_1 = \beta_2$ is tested. The chosen prior beliefs are the smallest values of a_0 and b_0 for which this hypothesis is not rejected, corresponding to the most uninformative prior. The motivation for this procedure is the inverse relationship between the dispersion of the prior beliefs and the magnitude of the option value. With more dispersed beliefs, the option value is greater, and the model needs to load less on option value to explain the investment paths, reducing the estimate of β_2 . The resulting inverse relationship

between the dispersion of the beliefs and β_2 together with the restriction $\beta_1 = \beta_2$ provides a way to estimate the relative dispersion. The parameter choice with the smallest values of a_0 and b_0 gives the most uninformative prior and lets the data shape the beliefs to a greater extent.

Still, this choice implies an initial value of λ_0 of 5% which is low relative to the empirical rate. This reflects a low expected return from investments in new technologies, and results in less exploration than under more optimistic prior beliefs. It should be noted that the estimates are sensitive to the choice of prior, and the low value of λ_0 is robust across specifications and, for example, changing the definition of success to only include IPOs (not acquisitions). There are several possible explanations for the low value of λ_0 . VCs may explore less than predicted by the model, either because they are irrational or because of informational frictions, leading to free-rider and herding problems and reducing the investors ability to internalize the value of information.¹⁴ Alternatively, investors may be specialized. This implies that they have lower success probabilities when investing outside their area of expertise, as captured by λ_0 , which measures expected success across all possible industry choices. Related, it is consistent with learning taking place at finer levels of individual technologies or ideas. If investors experiment within industries, and a failure in one technology leads them to try other investments within the same industry, the model would capture this reluctance to

¹⁴ See i.e. Amador and Weill (2006), Bolton and Harris (1999), Keller, Rady and Cripps (2005), Morris and Shin (2002).

abandon an industry, even after a string of failures, as being consistent with a low value of λ_0 . Finally, beliefs are may be misspecified. Assuming that investors have identical beliefs across all industries is probably unreasonably, but it is necessary for the empirical implementation, making it difficult to interpret the value of λ_0 literally. This problem is aggravated by the fact that VCs' histories prior to the sample period are unobserved, leading to the "initial conditions problem" discussed by Erdem and Keane (1996). To provide some evidence of this problem, the sample is broken into investors' early and late investments and the model is estimated separately for these two sub-samples. Estimates using VCs' initial 10 investments (the early sample) and the later investments (the late sample) are presented in the last specifications in Table II. For the second set of estimates the 10 initial investments are used to "burn-in" the VCs' beliefs and the beliefs should be more accurate for this sub-sample. The estimated coefficients are found to be more reasonable for the late sample, confirming that it is difficult to interpret the initial beliefs literally.

V. Summary and Conclusion

This paper demonstrates that VCs learn from past investments and anticipate to learn from future ones. When the distributions of payoffs from investments are uncertain, but the outcomes are informative about these distributions, the return from an investment consists of both its immediate return and an option value of learning. To empirically test for the presence of learning, the paper presents and estimates a learning model based on the Multi-armed Bandit model. It is shown that the *index result*

dramatically simplifies the estimation of this model, and the empirical approach based on this result may be applicable for investigating learning in other situations.

The empirical evidence confirms that learning is important, and the two alternative hypotheses, that (1) VCs do not learn, and (2) VCs learn only from past investments and do not internalize the value of future learning are both rejected. VCs exhibit *exploitative* behavior by changing their investments in response to the outcomes of past investments to benefit from higher immediate returns. Further, VCs exhibit *explorative* behavior by directing capital towards new unproven investments and internalizing the option value of the information gained from these investments. The model's further predictions are confirmed empirically. In the cross-section, VCs with more exploratory investment strategies have greater success rates, and VCs making more random investments are less successful. Further, more valuable investments are found to be made quicker. These findings are consistent with the model, and corroborate the economic interpretation of the measures of the immediate return and option value of learning derived under the model.

The presence of learning may have implications for understanding the market for entrepreneurial finance and the organization of VC firms. It violates the idea that individual investments are evaluated in isolation, and a growing theoretical literature shows that learning can lead to informational frictions. The two main ones are a free-rider problem and an informational cascading problem. Both of these frictions diminish investors' incentives to explore, reducing the equilibrium level of learning below the first-best level. However, VCs are private investors in privately held companies. This

organizational form reduces informational spillovers and may allow VCs to internalize the value of exploration and promote learning to a greater extent than traditional investors in more transparent capital markets. This ability to allocate capital to more explorative investments may present an additional source of value created by these investors.

One potential refinement of the analysis is to separate learning at the industry level from learning at the finer levels of the individual technology or business idea. The analysis is consistent with learning taking place at all of these levels, but an explicit separation of these various kinds of learning would sharpen the understanding of the learning process and its implications for the market. The main challenges are data availability and computational issues. With a finer classification of learning, there would be fewer investments within each category. Since the empirical inference is based on comparisons of investments in categories with different lengths of investment histories, longer overall histories would be required for sufficient variation. In addition, narrower classifications may make investments in one category informative about investments in related categories, further complicating the analysis.

Finally, Gompers, Kovner, Lerner and Scharfstein (2005) demonstrate that changes in economic fundamentals, as signaled by public market signals, are important determinants of VCs' investments. The analysis here may suggest that there are two kinds of changes in economic fundamentals that can make new investments attractive. An investment becomes attractive when either its immediate return increases (an upward shift in $F(t)$) or when its option value increases (an increase in the "spread" of $F(t)$). An increase in the immediate return may follow from an improvement of a known product,

for example through an investment to reduce its marginal cost. This would generate an immediate return, but it may have a small option value of learning. In contrast, an investment in a new and unproven technology typically has a low immediate return but a substantial option value, since a successful implementation of the technology would spur further investments in other applications. While standard capital markets are well suited for allocating capital in response to the first kind of changes, VCs may be better able to internalize the value of learning, making them better suited for allocating capital in response to the second kind of change. In fact, VCs invest primarily in entrepreneurial companies with new technologies and high option values, and formalizing the difference between standard capital markets and institutional investors in their abilities to internalize value and allocate capital may help understand the different roles of these financial institutions.

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Figure I:

The figure illustrates Option Value as a function of Lambda, keeping the number of investments constant. Option values are plotted for N equal to 40, 30, and 20, respectively.

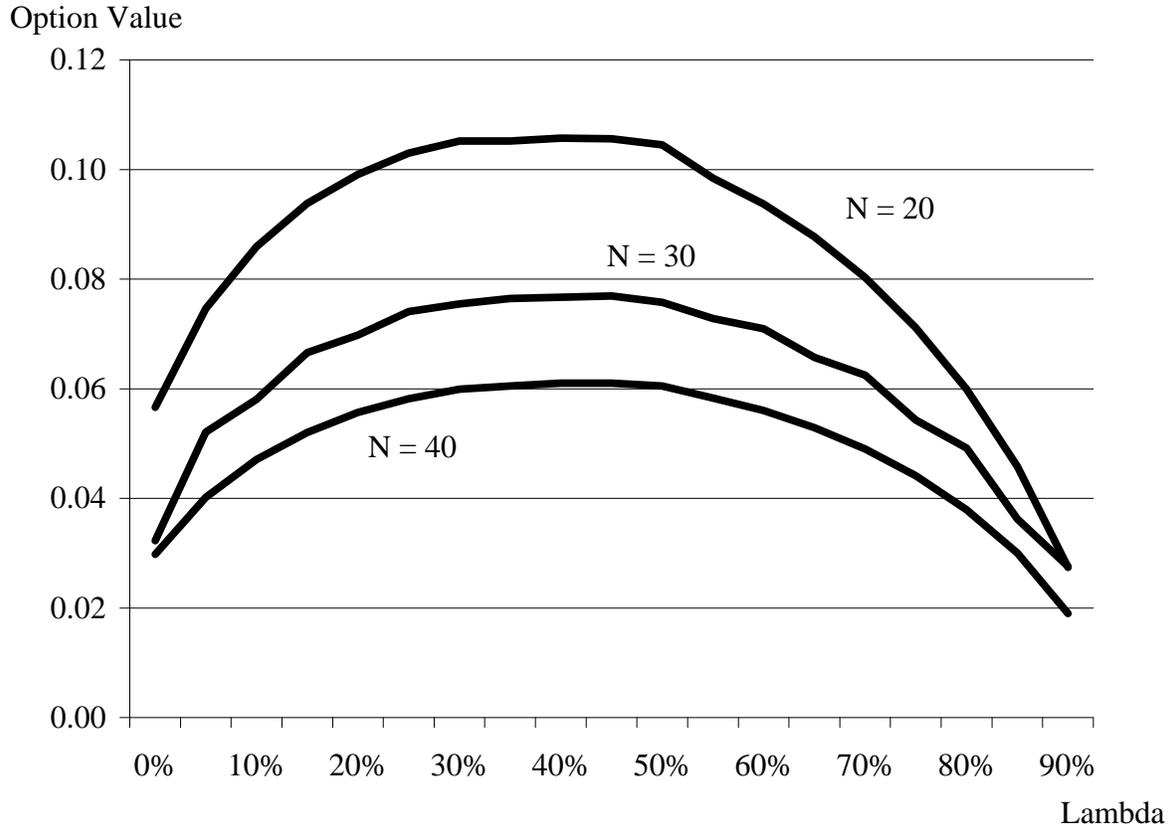


TABLE I: Summary Statistics

The table presents summary statistics of the sample. *IPO*, *Acquisition*, and *Success* are the fraction of companies going public, being acquired, and being classified as successful (IPO + Acquisition), respectively. *Year* is the year of the investment. Panel A presents characteristics at the company level. In Panel B characteristics are presented at the VC level, and the classifications of investment strategies are estimated as described in the text (normalized classifications have std. dev. equal to one in the sample). Panel C presents investment level characteristics. Values of *Lambda*, *Option Value*, *Gittins Index*, and *Option Value / Gittins Index* are calculated for the investments actually made. Across all potential choices, the means of these variables are 0.208, 0.062, 0.270, and 0.311, respectively. Panel D presents investments and IPOs per industry per year for the entire sample, included investors making less than 40 investments.

PANEL A: Summary Statistics By Company

	Obs.	Mean	Std. Dev.	Min	Max
IPO	6,076	0.170	0.376	0	1
Acquisition	6,076	0.329	0.470	0	1
Success (IPO + Acq)	6,076	0.499	0.500	0	1
Year	6,076	1995.2	4.422	1987	2000
Industry Classifications					
Health	6,076	0.181	0.385	0	1
Communciations	6,076	0.208	0.406	0	1
Computers	6,076	0.240	0.427	0	1
Software	6,076	0.175	0.380	0	1
Consumer	6,076	0.122	0.327	0	1
Other	6,076	0.074	0.261	0	1

PANEL B: Summary Statistics by Investor

	Obs.	Mean	Std. Dev.	Min	Max
IPO Rate	216	0.204	0.094	0.000	0.523
Acq Rate	216	0.299	0.067	0.106	0.492
Success Rate	216	0.503	0.118	0.133	0.864
Total Experience	216	88.731	64.210	40	577
Classifications of investment strategy:					
Option Value	216	2.762	27.551	-80.801	226.649
Standard Deviation	216	0.292	0.608	-4.213	4.640
Industry IPOs	216	0.002	0.016	-0.172	0.056
Normalized scale:					
Option Value	216	0.100	1.000	-2.933	8.226
Standard Deviation	216	0.480	1.000	-6.933	7.634
Industry IPOs	216	0.115	1.000	-11.040	3.602

TABLE I: Summary Statistics (cont.)**PANEL C: Summary Statistics by Investment**

	Obs.	Mean	Std. Dev.	Min	Max
Experience	19,166	68.081	73.230	1	577
Stage	19,166	0.287	0.452	0	1
Year	19,166	1995.2	4.542	1987	2000
Success	19,166	0.571	0.495	0	1
Lambda	19,166	0.260	0.153	0.032	0.717
Option Value	19,166	0.062	0.012	0.019	0.082
Gittins Index	19,166	0.323	0.147	0.063	0.745
OptionValue / Gittins Index	19,166	0.252	0.138	0.031	0.531

PANEL D: INVESTMENTS (IPOs) PER INDUSTRY PER YEAR

Year	Health	Comm	Comp	Cons	Soft	Other	Total
1987	806 (3)	420 (2)	1,125 (4)	164 (0)	359 (0)	310 (2)	3,184 (11)
1988	592 (4)	237 (4)	770 (5)	92 (1)	327 (1)	208 (4)	2,226 (19)
1989	395 (11)	148 (1)	371 (7)	57 (4)	189 (4)	139 (5)	1,299 (32)
1990	283 (11)	101 (4)	256 (10)	56 (1)	196 (6)	89 (4)	981 (36)
1991	258 (45)	100 (11)	164 (14)	57 (2)	206 (8)	54 (3)	839 (83)
1992	372 (59)	152 (15)	142 (18)	44 (11)	257 (11)	58 (6)	1,025 (120)
1993	357 (35)	159 (14)	123 (36)	74 (9)	163 (18)	57 (15)	933 (127)
1994	338 (33)	176 (13)	159 (27)	65 (6)	189 (16)	36 (6)	963 (101)
1995	445 (40)	240 (17)	191 (30)	134 (5)	280 (33)	86 (6)	1,376 (131)
1996	472 (72)	429 (35)	291 (28)	176 (16)	458 (43)	93 (13)	1,919 (207)
1997	621 (39)	533 (18)	391 (21)	269 (9)	633 (17)	124 (10)	2,571 (114)
1998	688 (9)	723 (23)	426 (16)	410 (9)	739 (11)	231 (2)	3,217 (70)
1999	844 (14)	1,938 (94)	1,055 (27)	1714 (53)	1,249 (68)	175 (3)	6,975 (259)
2000	961 (60)	2,824 (44)	3,443 (29)	1,388 (32)	1,034 (48)	359 (8)	10,009 (221)
Total	7,432 (435)	8,180 (295)	8,907 (272)	4,700 (158)	6,279 (284)	2,019 (87)	37,517 (1,531)

TABLE II: Investment Decisions

The table reports estimates of a Multinomial Logit model (McFadden choice model) where investors' industry choice is the endogenous variable. The possible choices are Health, Communications, Computers, Consumer Goods, Software, and Other. *Lambda* and *Option Value* are investors' expected immediate return and option value of investing. *Industry Investments* is total number of investments in each industry per year across all investors in the data. *Industry Experience* is the past number of investments by the investor in the industry. *Previous* is a binary variable that equals one for the industry of the investor's previous investment. Early Sample and Late Sample are estimated using each investors initial 10 investments, and investments 11 and up, respectively. Robust standard errors with clustering at the company level are in parenthesis. ***, **, and * denote statistical significance at the 1%, 5%, and 10% levels, respectively.

	1	2	3	4	5	Early Sample	Late Sample
Lambda	4.7657 *** (.0718)	4.2445 *** (.0788)	4.3202 *** (.0780)	3.3410 *** (.1354)	2.4023 *** (.1484)	17.8877 *** (1.7681)	4.6238 *** (.1008)
Option Value	6.4545 *** (.8884)	5.3043 *** (.9006)	3.0443 *** (.9344)	17.2952 *** (1.2284)	8.5117 *** (1.3057)	-33.0524 *** (8.7047)	3.7897 *** (1.0749)
Industry IPOs		0.0085 *** (.0004)			0.0188 *** (.0018)		
Industry Investments			0.0007 *** (.0000)		0.0006 *** (.0000)		
Industry Experience				0.0228 *** (.0018)	0.0166 *** (.0019)		
Previous					0.2978 *** (.0174)		
Industry Controls	No	No	No	No	Yes	No	No
Observations	19,166	19,166	19,166	19,166	19,166	2,160	17,006

TABLE III: Investment Outcomes

The table reports marginal effect from estimates of a Probit model where the outcome (success or failure) of each investment is the endogenous variable. *Lambda* and *Option Value* are investors' expected immediate return and option value of investing. *Industry Investments* is total number of investments in each industry per year across all investors in the data. *Industry IPOs* is the number of companies in the same industry going public in the year of the investment. *Industry Experience* is the past number of investments by the investor in the industry and *Total Experience* is total number of the investor's past investments across all industries. In the third specification, investors initial ten investments are discarded from burn-in. Robust standard errors with clustering at the company level are in parenthesis. ***, **, and * denote statistical significance at the 1%, 5%, and 10% levels, respectively.

	1	2	3
Lambda	0.3631 *** (0.0368)	0.3544 *** (0.0541)	0.6668 *** (0.0772)
Option Value	3.1726 *** (0.3975)	1.0309 * (0.5772)	0.5607 (0.6615)
Stage	0.1656 *** (0.0141)	0.1601 *** (0.0142)	0.1726 *** (0.0147)
Industry Experience		-0.0032 *** (0.0008)	-0.0037 *** (0.0098)
Total Experience		0.0004 *** (0.0001)	0.0008 *** (0.0002)
Industry IPOs			-0.0004 (0.0004)
Industry Investments			0.0000 (0.0000)
log(Industry Experience +1)			-0.0363 ** (0.0159)
log(Total Experience +1)			-0.0397 *** (0.0159)
Year Controls	Yes	Yes	Yes
Industry Controls	No	Yes	Yes
Observations	19,166	19,166	17,006

TABLE IV: Investment Strategies and Outcomes

Panel A shows estimated coefficients for an OLS regression. An observation is an investor and the endogenous variable is the investor's success rate. Panel B presents marginal effects estimated from a Probit model. Each observation is an investment in a company and the endogenous variable is the outcome. *Option Value*, *Standard Error*, and *Industry IPOs* characterize the investor's investment strategy in terms of its dependence on option value, its standard error, and on the number of VC backed IPOs in the industry in the same year. These coefficients are normalized to have standard error equal one (see text for details). *Total Experience* measures the number of previous investments by the investor at the time of each investment. *Final Experience* is the investor's experience at the end of the sample. *Stage* is an indicator variable that equals one for investments in late-stage companies. Observations are weighted according to the precision of the estimates (see text for details). Robust standard errors with clustering at the company level are in parenthesis. ***, **, and * denote statistical significance at the 1%, 5%, and 10% levels, respectively.

PANEL A: Success Rate of Venture Capital Firms									
	1			2			3		
	Coef.	Std. Err.		Coef.	Std. Err.		Coef.	Std. Err.	
Classification of Strategy									
Option Value	0.0214	(.0027)	***	0.0234	(.0087)	***	0.0262	(.0089)	***
Standard Deviation	-0.0160	(.0022)	***	-0.0212	(.0071)	***	-0.0240	(.0080)	***
Industry IPOs				0.0085	(.0112)		0.0086	(.0117)	
Final Experience							0.0001	(.0001)	
Constant	0.5432	(.0044)	***	0.5441	(.0094)	***	0.5285	(.0126)	***
Observations	216			216			216		
PANEL B: Success of Individual Investments									
	1			2			3		
	dF/dX	Std. Err.		dF/dX	Std. Err.		dF/dX	Std. Err.	
Classification of Strategy									
Option Value	0.0335	(.0058)	***	0.0148	(.0060)	***	0.0154	(.0060)	***
Standard Deviation	-0.0299	(.0062)	***	-0.0242	(.0075)	***	-0.0212	(.0076)	***
Industry IPOs				0.0029	(.0094)		-0.0013	(.0095)	
Stage							0.1514	(.0157)	***
Total Experience							0.0001	(.0001)	
Industry IPOs							-0.0002	(.0005)	
Year Controls	No			Yes			Yes		
Industry Controls	No			No			Yes		
Observations	19,166			19,166			19,166		

TABLE V: Investment Speed

The table reports estimated coefficients from four OLS regressions and two Cox Hazard models. An observation is an investment, and the time since the previous investment (measured in days) is the endogenous variable. Each investor's initial investment and investments made less than 14 days apart are excluded. *Gittins* is the Gittins index of the investment, *Option Value* is the option value, and *Lambda* is the expected immediate return. *Total Experience* is the investor's experience, measured as the total number of past investments. Standard errors clustered at the company level are reported in parenthesis. ***, **, and * denote statistical significance at the 1%, 5%, and 10% levels, respectively.

	OLS				Cox Hazard	
	1	2	3	4	5	6
Gittins	-119.85 *** (9.00)	-119.54 *** (13.06)			3.88 *** (0.28)	
Option Value			385.70 *** (118.47)	-822.23 *** (149.28)		34.23 *** (36.86)
Lambda			-108.55 *** (10.30)	-116.93 *** (12.87)		2.35 *** (0.26)
Total Experience		-0.35 *** (0.03)		-0.45 *** (0.04)		1.01 *** (0.00)
Constant	118.12 *** (3.68)	128.80 *** (5.84)	82.75 *** (9.90)	174.50 *** (12.47)		
Industry Controls	No	Yes	No	Yes	No	Yes
Year Controls	No	Yes	No	Yes	No	Yes
Observations	10,881	10,881	10,881	10,881	10,881	10,881